



# Quantitative functional failure analysis of a thermal-hydraulic passive system by means of bootstrapped Artificial Neural Networks

Enrico Zio, George Apostolakis, Nicola Pedroni

## ► To cite this version:

Enrico Zio, George Apostolakis, Nicola Pedroni. Quantitative functional failure analysis of a thermal-hydraulic passive system by means of bootstrapped Artificial Neural Networks. *Annals of Nuclear Energy*, Elsevier Masson, 2010, 37 (5), pp.639-649. .

**HAL Id: hal-00609176**

**<https://hal-supelec.archives-ouvertes.fr/hal-00609176>**

Submitted on 26 Jul 2012

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

# Quantitative functional failure analysis of a thermal-hydraulic passive system by means of bootstrapped Artificial Neural Networks

E. Zio<sup>\*a</sup>, G. E. Apostolakis<sup>b</sup>, N. Pedroni<sup>a</sup>

<sup>a</sup>Energy Department, Politecnico di Milano, Via Ponzio 34/3, 20133 Milan, Italy

\*Phone: +39-2-2399-6340; fax: +39-2-2399-6309

E-mail address: enrico.zio@polimi.it

<sup>b</sup>Department of Nuclear Science and Engineering, Massachusetts Institute of Technology,

77 Massachusetts Avenue, Cambridge (MA) 02139-4307

Phone: +1-617-252-1570; fax: +1-617-258-8863

E-mail address: apostola@mit.edu

## Abstract

*The estimation of the functional failure probability of a thermal-hydraulic (T-H) passive system can be done by Monte Carlo (MC) sampling of the epistemic uncertainties affecting the system model and the numerical values of its parameters, followed by the computation of the system response by a mechanistic T-H code, for each sample. The computational effort associated to this approach can be prohibitive because a large number of lengthy T-H code simulations must be performed (one for each sample) for accurate quantification of the functional failure probability and the related statistics.*

*In this paper, the computational burden is reduced by replacing the long-running, original T-H code by a fast-running, empirical regression model: in particular, an Artificial Neural Network (ANN) model is considered. It is constructed on the basis of a limited-size set of data representing examples of the input/output nonlinear relationships underlying the original T-H code; once the model is built, it is used for performing, in an acceptable computational time, the numerous system response calculations needed for an accurate failure probability estimation, uncertainty propagation and sensitivity analysis.*

*The empirical approximation of the system response provided by the ANN model introduces an additional source of (model) uncertainty, which needs to be evaluated and accounted for. A bootstrapped ensemble of ANN regression models is here built for quantifying, in terms of confidence intervals, the (model) uncertainties associated with the estimates provided by the ANNs. For demonstration purposes, an application to the functional failure analysis of an emergency passive decay heat removal system in a simple steady-state model of a Gas-cooled Fast Reactor*

(GFR) is presented. The functional failure probability of the system is estimated together with global Sobol sensitivity indices. The bootstrapped ANN regression model built with low computational time on few (e.g., 100) data examples is shown capable of providing reliable (very near to the true values of the quantities of interest) and robust (the confidence intervals are satisfactorily narrow around the true values of the quantities of interest) point estimates.

## 1 Introduction

Passive systems (IAEA, 1991) are expected to contribute significantly to the safety of future nuclear power plants by combining their peculiar characteristics of simplicity, reduction of human interaction (Prosek and Cepin, 2008) and reduction or avoidance of hardware failures (Mathews et al., 2008). On the other hand, the uncertainties associated to their actual operation and modelling are usually larger than in active systems.

Two different sources of uncertainties are usually considered in safety analyses: randomness due to intrinsic variability in the behavior of the system (aleatory uncertainty) and imprecision due to lack of data on some underlying phenomena (e.g., natural circulation) and to scarce or null operating experience over the wide range of conditions encountered during operation (epistemic uncertainty) (Apostolakis, 1990; Helton, 2004).

Due to these uncertainties, there is a nonzero probability that the physical phenomena involved in the operation of a passive system do not occur as expected, thus leading to the failure of performing the intended safety function even if safety margins have been dimensioned (Marquès et al., 2005; Patalano et al., 2008).

Various methodologies have been proposed in the open literature to quantify the probability that T-H passive systems fail to perform their functions: the reader is referred to (Burgazzi, 2007; Zio and Pedroni, 2009a) for a review. A reasonable approach to take is that founded on the concept of *functional failures* in the framework of reliability physics and load-capacity exceedance probability (Burgazzi, 2003). In this view, a passive system fails to perform its function due to deviations from its expected behavior which lead the load imposed on the system (e.g., the peak value of the fuel cladding temperature during a LOCA transient) to overcome its capacity (e.g., a threshold value imposed by regulating authorities or by the mechanical properties of structural materials). This concept is at the basis of many works of literature, including (Jafari et al., 2003; Marquès et al., 2005; Pagani et al., 2005; Bassi and Marquès, 2008; Mackay et al., 2008; Mathews et al., 2008 and 2009; Patalano et al., 2008; Fong et al., 2009; Zio and Pedroni, 2009a and b); in these works, the passive system is modeled by a detailed, mechanistic T-H system code and the probability of not performing the required function is estimated based on a Monte Carlo (MC) sample of code runs

which propagate the *epistemic* (state-of-knowledge) uncertainties in the model describing the system and the numerical values of its parameters/variables; because of these uncertainties, the system may not accomplish its mission, even if no hardware failure occurs.

MC simulation allows a realistic assessment of the T-H system functional failure probability, thanks to its flexibility and indifference to the complexity of the T-H model. This, however, is paid in terms of the high computational efforts required. Indeed, a large number of Monte Carlo-sampled T-H model evaluations must generally be carried out for an accurate estimation of the functional failure probability. Since the number of simulations required to obtain a given accuracy depends on the magnitude of the failure probability to be estimated, with the computational burden increasing with decreasing functional failure probability (Schueller, 2007), this poses a significant challenge for the typically quite small (e.g., less than  $10^{-4}$ ) probabilities of functional failure of T-H passive safety systems. In particular, the challenge is due to the time required for each run of the detailed, mechanistic T-H system code which can take several hours (one code run is required for each sample of values drawn from the uncertainty distributions of the input parameters/variables to the T-H code) (Fong et al., 2009). Thus, alternative methods must be sought to tackle the computational burden associated to the analysis.

To tackle the computational issue, efficient sampling techniques can be adopted for obtaining robust estimations with a limited number of input samples. Techniques like Importance Sampling (IS) (Au and Beck, 2003; Schueller et al., 2004), Stratified Sampling (Helton and Davis, 2003; Cacuci and Ionescu-Bujor, 2004) and Latin Hypercube Sampling (LHS) (Helton and Davis, 2003) have been widely used in reliability analysis and risk assessment (Helton 1998). Recently, advanced sampling methods such as Subset Simulation (SS) (Au and Beck, 2001; Au and Beck, 2003) and Line Sampling (LS) (Koutsporelakis et al., 2004; Pradlwarter et al., 2005) have been proposed for structural reliability assessment and subsequently applied to the estimation of the functional failure probability of a T-H passive system (Zio and Pedroni, 2009a and b)<sup>1</sup>. These methods have been shown to improve the computational efficiency although there is no indication yet that the number of model evaluations can be reduced to below a few hundreds, which may be mandatory in the presence of computer codes requiring several hours to run a single simulation.

---

<sup>1</sup> Apart from efficient MC techniques, there exist methods based on nonparametric order statistics (Wilks, 1942) that propagate uncertainties through mechanistic T-H codes with reduced computational burden, especially if only one- or two-sided confidence intervals are needed for particular statistics (e.g., the 95<sup>th</sup> percentile) of the outputs of the code. For example, the so-called *coverage* (Guba et al., 2003; Makai and Pal, 2006) and *bracketing* (Nutt and Wallis, 2004) approaches can be used to identify the *precise* number of sample code runs required to obtain a given *confidence* level on the estimates of prescribed statistics of the code outputs.

In such cases, the only viable alternative seems that of resorting to fast-running, surrogate regression models, also called response surfaces or meta-models, to approximate the input/output function implemented in the long-running T-H model code, and then substitute it in the passive system functional failure analysis. The construction of such regression models entails running the T-H model code a predetermined, *reduced* number of times (e.g., 50-100) for specified values of the uncertain input parameters/variables and collecting the corresponding values of the output of interest; then, statistical techniques are employed for fitting the response surface of the regression model to the input/output data generated in the previous step. Several examples can be found in the open literature concerning the application of surrogate meta-models in reliability problems. In (Bucher and Most, 2008; Gavin and Yau, 2008; Liel et al., 2008), polynomial Response Surfaces (RSs) are employed to evaluate the failure probability of structural systems; in (Arul et al., 2009; Fong et al., 2009; Mathews et al., 2009), *linear* and *quadratic* polynomial RSs are employed for performing the reliability analysis of T-H passive systems in advanced nuclear reactors; in (Deng, 2006; Hurtado, 2007; Cardoso et al., 2008; Cheng et al., 2008), learning statistical models such as Artificial Neural Networks (ANNs), Radial Basis Functions (RBFs) and Support Vector Machines (SVMs) are trained to provide local approximations of the failure domain in structural reliability problems; in (Volkova et al., 2008; Marrel et al., 2009), Gaussian meta-models are built to calculate global sensitivity indices for a complex hydrogeological model simulating radionuclide transport in groundwater.

In this work, Artificial Neural Networks (ANNs) are considered to reduce the computational burden associated to uncertainty propagation in the functional failure analysis of a natural convection-based decay heat removal system of a Gas-cooled Fast Reactor (GFR) (Pagani et al., 2005). A *limited-size* set of input/output data examples is used to construct the ANN regression model; once the model is built, it is used to perform, in a negligible computational time, the functional failure analysis of the T-H passive system: in particular, the functional failure probability of the system is estimated together with global sensitivity indices of the naturally circulating coolant temperature.

The use of regression models in safety critical applications like nuclear power plants raises concerns with regards to the *model* accuracy, which must be not only verified but also quantified; in this paper, we resort to the bootstrap method for quantifying, in terms of *confidence intervals*, the model uncertainty associated to the estimates provided by ANNs (Efron and Tibshirani, 1993). Some examples of the application of the bootstrap method for the evaluation of the uncertainties associated to the output of regression models in safety-related problems can be found in the literature: in (Zio, 2006), bootstrapped ANNs are trained to predict nuclear transients processes; in

(Cadini et al., 2008; Secchi et al., 2008), the model uncertainty, quantified in terms of a *standard deviation*, is used to “correct” the ANN output in order to provide conservative estimates for important safety parameters in nuclear reactors (i.e., percentiles of the pellet cladding temperature); finally, in (Storlie et al., 2008), the bootstrap procedure is combined with different regression techniques, e.g. Multivariate Adaptive Regression Spline (MARS), Random Forest (RF) and Gradient Boosting Regression (GBR), to calculate confidence intervals for global sensitivity indices of the computationally demanding model of a nuclear waste repository.

The main objective of the present study is to show the possibility of:

- limiting the computational burden associated to the uncertainty propagation and sensitivity analyses underpinning a quantitative functional failure analysis of a T-H passive system; to this aim, ANN regression models trained on a small data set are used for the first time for the estimation of:
  - the functional failure probability of the passive system;
  - first-order global Sobol sensitivity indices determining the contribution of the individual uncertain parameters (i.e., the inputs to the T-H code) to the uncertainty in the coolant temperature (i.e., the output of the T-H code) and consequently to the probability of functional failure of the passive system;
- quantifying, in terms of *confidence intervals*, the model uncertainty associated to the estimates of the functional failure probability and Sobol indices provided by the ANN models. To the best of the authors’ knowledge, the issue of quantification of model uncertainty in the regression estimates has not been addressed in the literature regarding the functional failure analysis of T-H passive systems (Arul et al., 2009; Fong et al., 2009; Mathews et al., 2009); in the present work, bootstrapped ANNs are used to the purpose.

The paper organization is as follows. In Section 2, the concepts of functional failure analysis for T-H passive systems are synthetically summarized. Section 3 briefly presents the problem of empirical regression modeling by means of ANNs and provides a snapshot on the bootstrap method for the quantification of the ANN (model) uncertainty. In Section 4, the case study of literature concerning the passive cooling of a GFR is presented. In Section 5, the results of the application of bootstrapped ANNs to the functional failure analysis of the T-H passive system of Section 4 are reported. Conclusions are provided in the last section. Finally, algorithmic details about the bootstrap-based method for quantifying, in terms of *confidence intervals*, the model uncertainty

associated to the estimates of safety parameters computed by ANN regression models are reported in the Appendix at the end of the paper.

## 2 Functional failure analysis of T-H passive systems

Since a comprehensive functional failure analysis of a T-H passive system is beyond the scope of this work, only the essential steps for the conceptual development of the analysis are briefly reported below (Marquès et al., 2005):

1. Define the failure criteria of the passive system.
2. Build a deterministic, best-estimate T-H model to simulate the passive system behavior.
3. Identify and characterize by proper probability distributions the factors of *uncertainty* in the results of the best estimate T-H calculations. In the present work, only *epistemic* uncertainties are considered, due to the limited knowledge on some phenomena and processes (e.g., models, parameters and correlations used in the T-H analysis).
4. Propagate by Monte Carlo Simulation (MCS), the epistemic uncertainties associated to the identified relevant parameters, models and correlations (step 3. above) through the deterministic, long-running T-H code in order to estimate the functional failure probability of the passive system conditional on the current state of knowledge about the phenomena involved (step 3. above) (Bassi and Marques, 2008; Mackay et al., 2008; Mathews et al., 2008; Patalano et al., 2008). Formally, let  $\mathbf{x} = \{x_1, x_2, \dots, x_j, \dots, x_{n_i}\}$  be the vector of the relevant uncertain system parameters/variables,  $Y(\mathbf{x})$  be a scalar indicator variable of the performance of the passive system (e.g., the fuel peak cladding temperature) and  $\alpha_Y$  a threshold value defining the corresponding failure criterion (e.g., a limit value imposed by regulating authorities). For illustrating purposes, let us assume that the passive system operates as long as  $Y(\mathbf{x}) < \alpha_Y$ . The MCS procedure for estimating the functional failure probability entails that a *large* number  $N_T$  of samples of the values of the system parameters  $\mathbf{x}$  be drawn from the corresponding probability distributions and used to evaluate  $Y(\mathbf{x})$  by running the T-H code. An estimate  $\hat{P}(F)$  of the probability of failure  $P(F)$  can then be computed by dividing the number of times that  $Y(\mathbf{x}) > \alpha_Y$  by the total number of samples  $N_T$ .
5. Perform a sensitivity study to determine the contribution of the individual uncertain parameters (i.e., the inputs to the T-H code) to the uncertainty in the outputs of the T-H code and consequently to the functional failure probability of the T-H passive system. As is true for uncertainty propagation (step 4. above), sensitivity analysis relies on *multiple* evaluations of the code for different combinations of system inputs.

The computational burden posed by the uncertainty propagation and sensitivity analysis of steps 4. and 5. above may be tackled by replacing the long-running, original T-H model code by a fast-running, surrogate regression model properly built to approximate the output from the true system model. In this paper, Artificial Neural Networks (ANNs) (Bishop, 1995) are used. A confidence interval is evaluated by means of the bootstrap method (Efron and Tibshirani, 1993); a brief description of this latter technique is given in the following Section, whereas the relevant algorithmic details can be found in the Appendix at the end of the paper.

### 3 Bootstrapped Artificial Neural Networks for empirical regression modelling

As discussed in the previous Section, the computational burden posed by uncertainty and sensitivity analyses of T-H passive systems can be tackled by replacing the long-running, original T-H model code by a fast-running, surrogate regression model. Because calculations with the surrogate model can be performed quickly, the problem of long simulation times is circumvented.

Let us consider a generic meta-model to be built for performing the task of nonlinear regression, i.e., estimating the nonlinear relationship between a vector of input variables  $\mathbf{x} = \{x_1, x_2, \dots, x_j, \dots, x_{n_i}\}$  and a vector of output targets  $\mathbf{y} = \{y_1, y_2, \dots, y_l, \dots, y_{n_o}\}$ , on the basis of a *finite* (and possibly *reduced*) set of input/output data examples (i.e., patterns),  $D_{train} = \{\{\mathbf{x}_p, \mathbf{y}_p\}, p=1, 2, \dots, N_{train}\}$  (Zio, 2006). It can be assumed that the target vector  $\mathbf{y}$  is related to the input vector  $\mathbf{x}$  by an unknown nonlinear deterministic function  $\boldsymbol{\mu}_y(\mathbf{x})$  corrupted by a noise vector  $\boldsymbol{\varepsilon}(\mathbf{x})$ , i.e.,

$$\mathbf{y}(\mathbf{x}) = \boldsymbol{\mu}_y(\mathbf{x}) + \boldsymbol{\varepsilon}(\mathbf{x}). \quad (1)$$

Notice that in the present case of T-H passive system functional failure probability assessment the vector  $\mathbf{x}$  contains the relevant uncertain system parameters/variables, the nonlinear deterministic function  $\boldsymbol{\mu}_y(\mathbf{x})$  represents the complex, long-running T-H mechanistic model code (e.g., RELAP5-3D), the vector  $\mathbf{y}(\mathbf{x})$  contains the output variables of interest for the analysis and the noise  $\boldsymbol{\varepsilon}(\mathbf{x})$  represents the errors introduced by the numerical methods employed to calculate  $\boldsymbol{\mu}_y(\mathbf{x})$  (Storlie et al., 2008); for simplicity, in the following we assume  $\boldsymbol{\varepsilon}(\mathbf{x}) = \mathbf{0}$  (Secchi et al., 2008).

The objective of the regression task is to estimate  $\boldsymbol{\mu}_y(\mathbf{x})$  in (1) by means of a regression function  $f(\mathbf{x}, \mathbf{w}^*)$  depending on a set of parameters  $\mathbf{w}^*$  to be properly determined on the basis of the available data set  $D_{train}$ ; the algorithm used to calibrate the set of parameters  $\mathbf{w}^*$  is obviously dependent on the nature of the regression model adopted, but in general it aims at minimizing the mean (absolute or



quadratic) error between the output targets of the original T-H code,  $y_p = \boldsymbol{\mu}_y(\mathbf{x}_p)$ ,  $p = 1, 2, \dots, N_{train}$ , and the output vectors of the regression model,  $\hat{y}_p = \mathbf{f}(\mathbf{x}_p, \mathbf{w}^*)$ ,  $p = 1, 2, \dots, N_{train}$ ; for example, the Root Mean Squared Error (RMSE) is commonly adopted to this purpose (Zio, 2006):

$$RMSE = \frac{1}{N_{train} \cdot n_o} \sum_{p=1}^{N_{train}} \sum_{l=1}^{n_o} (y_{p,l} - \hat{y}_{p,l})^2. \quad (2)$$

Once built, the regression model  $\mathbf{f}(\mathbf{x}, \mathbf{w}^*)$  can be used as a simplified, quick-running surrogate of the original, long-running T-H model code for significantly reducing the computational burden associated to the accurate estimation and epistemic uncertainty propagation steps of the functional failure analysis of T-H passive systems. In particular, the regression model  $\mathbf{f}(\mathbf{x}, \mathbf{w}^*)$  can be used in place of the T-H code to calculate any quantity of interest  $Q$ , e.g. the functional failure probability of the passive system, confidence intervals and global sensitivity indices, etc.

In this work, three-layered feed-forward Artificial Neural Network (ANN) regression models are considered. In extreme synthesis, ANNs are computing devices inspired by the function of the nerve cells in the brain (Bishop, 1995). They are composed of many parallel computing units (called neurons or nodes) interconnected by weighed connections (called synapses). Each of these computing units performs a few simple operations and communicates the results to its neighbouring units. From a mathematical viewpoint, ANNs consist of a set of nonlinear (e.g., sigmoidal) basis functions with adaptable parameters  $\mathbf{w}^*$  that are adjusted by a process of *training* (on many different input/output data examples), i.e., an iterative process of regression error minimization (Rumelhart et al., 1986). The details about ANN regression models are not reported here for brevity, the interested reader may refer to the cited references and the copious literature in the field. The particular type of ANN employed in this paper is the classical three-layered feed-forward ANN trained by the error back-propagation algorithm.

When using the approximation of the system output provided by an ANN empirical regression model, an additional source of uncertainty is introduced which needs to be evaluated, particularly in safety critical applications like those related to nuclear power plant technology. One way to do this by resorting to bootstrapped ANN regression models (Efron and Tibshirani, 1993), i.e., an ensemble of ANN regression models, constructed on different data sets bootstrapped from the original one (Zio, 2006; Storlie et al., 2008). The bootstrap method is a distribution-free inference method which requires no prior knowledge about the distribution function of the underlying population (Efron and Tibshirani, 1993). The basic idea is to generate a sample from the observed data by sampling with replacement from the original data set (Efron and Tibshirani, 1993). From the theory and practice of ensemble empirical models, it can be shown that the estimates given by

bootstrapped ANN regression models is in general more accurate than the estimate of the best ANN regression model in the bootstrap ensemble of ANN regression models (Zio, 2006; Cadini et al., 2008).

A bootstrap-based estimation technique is here employed for the evaluation of the so-called Bootstrap Bias Corrected (BBC) point estimate  $\hat{Q}_{BBC}$  of a generic quantity  $Q$  (e.g., a safety parameter) by an ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  and the calculation of the associated BBC Confidence Interval (CI) (Zio, 2006; Storlie et al., 2008). The complete algorithm is not reported here for brevity; some details can be found in the Appendix at the end of the paper.

#### **4 Case study: functional failure analysis of a T-H passive system**

The case study considered in this work concerns the natural convection cooling in a Gas-cooled Fast Reactor (GFR) under a post-Loss Of Coolant Accident (LOCA) condition (Pagani et al., 2005). The reactor is a 600-MW GFR cooled by helium flowing through separate channels in a silicon carbide matrix core whose design has been the subject of study in the past several years at the Massachusetts Institute of Technology (MIT) (Pagani et al., 2005).

In these studies, the possibility of using natural circulation to remove the decay heat in case of an accident is demonstrated. In particular, in the case of a LOCA, the long-term heat removal is ensured by natural circulation in a given number  $N_{loops}$  of identical and parallel loops.

A GFR decay heat removal configuration is shown schematically in Figure 1; only one of the  $N_{loops}$  loops is reported for clarity of the picture: the flow path of the cooling helium gas is indicated by the black arrows; the loop has been divided into  $N_{sections} = 18$  sections for numerical calculation; technical details about the geometrical and structural properties of these sections are not reported here for brevity: the interested reader may refer to (Pagani et al., 2005).

**Figure 1 here**

In the present analysis, the average core power to be removed is assumed to be 18.7 MW, equivalent to about 3% of full reactor power (600 MW): to guarantee natural circulation cooling at this power level, a pressure of 1650 kPa is required in nominal conditions. Finally, the secondary side of the heat exchanger (i.e., item 12 in Figure 1) is assumed to have a nominal wall temperature of 90 °C (Pagani et al., 2005).

The model describes the quasi-steady-state natural circulation cooling that takes place after the LOCA transient has occurred. The associated simplifications introduced in the modeling allow relatively fast calculations, which enable to obtain reference values for comparison of the estimates

obtained by the ANN models developed. From a strictly mathematical point of view, obtaining a steady-state solution amounts to dropping the time dependent terms in the energy and momentum conservation equations. In practice, the T-H model code balances the pressure losses around the loops so that friction and form losses are compensated by the buoyancy term, while at the same time maintaining the heat balance in the heater (i.e., the reactor core, item 4 in Figure 1) and cooler (i.e., the heat exchanger, item 12 in Figure 1); a thorough description of the deterministic T-H model is not given here for brevity: the interested reader may refer to (Pagani et al., 2005) for details.

#### 4.1 Uncertainties in the T-H model

Uncertainties affect the actual operation of passive systems and its modeling. On the one side, there are phenomena, like the occurrence of unexpected events and accident scenarios, e.g. the failure of a component or the variation of the geometrical dimensions and material properties, which are random in nature. This kind of uncertainty in the model description of the system behavior is termed aleatory (NUREG-1150, 1990; Helton, 1998; USNRC, 2002). In this work, as well as in the reference paper by (Pagani et al., 2005), aleatory uncertainties are not considered for the estimation of the functional failure probability of the T-H passive system of Figure 1.

An additional contribution of uncertainty comes from the incomplete knowledge on the properties of the system and the conditions in which the phenomena occur (i.e., natural circulation). This uncertainty is often termed epistemic and affects the model representation of the system behaviour, in terms of both (*model*) uncertainty in the hypotheses assumed and (*parameter*) uncertainty in the values of the parameters of the model (Cacuci and Ionescu-Bujor, 2004; Helton et al., 2006; Patalano et al., 2008).

Model uncertainty arises because mathematical models are simplified representations of real systems and, therefore, their outcomes may be affected by errors or bias. It may, for example, involve the correlations adopted to describe the T-H phenomena, which are subject to errors of approximation. Such uncertainties may for example be represented by a multiplicative model (Zio and Apostolakis, 1996; Patalano et al., 2008):

$$z = m(\mathbf{x}) \cdot \zeta, \quad (3)$$

where  $z$  is the real value of the parameter to be correlated (e.g., heat transfer coefficients, friction factors, Nusselt numbers or thermal conductivity coefficients),  $m(\cdot)$  is the mathematical model of the correlation,  $\mathbf{x}$  is the vector of correlating variables and  $\zeta$  is a multiplicative error factor. Hence, the uncertainty in the output quantity  $z$  is translated into an uncertainty in the multiplicative error factor  $\zeta$ , commonly classified as representing *model* uncertainty.

Uncertainty affects also the values of the *parameters* used in the model (e.g., power level, pressure, cooler wall temperature, material conductivity, ...), e.g. owing to errors in their measurement or insufficient data and information. As a consequence, the values of such parameters are usually known only to a certain level of precision, i.e., epistemic uncertainty is associated with them (Pagani et al., 2005).

In this work, *only* epistemic (i.e., model and parameter) uncertainties are represented and propagated through the deterministic T-H code (Pagani et al., 2005; Bassi and Marques, 2008; Mackay et al., 2008; Mathews et al., 2008; Patalano et al., 2008). Parameter uncertainties are associated to the reactor power level, the pressure in the loops after the LOCA and the cooler wall temperature. Model uncertainties are associated to the correlations used to calculate the Nusselt numbers and friction factors in the forced, mixed and free convection regimes. The corresponding nine uncertain inputs of the model  $\{x_j : j = 1, 2, \dots, 9\}$  are assumed to be distributed according to normal distributions of known mean  $\mu$  and standard deviation  $\sigma$ , taking values in the range  $[\mu - 4\sigma, \mu + 4\sigma]$  (Table 1, Pagani et al., 2005). The practical and conceptual reasons underpinning the values in Table 1 are described in (Pagani et al., 2005).

**Table 1 here**

## 4.2 Failure criteria of the T-H passive system

The passive decay heat removal system of Figure 1 is considered failed whenever the temperature of the coolant helium leaving the core (item 4 in Figure 1) exceeds either 1200 °C in the hot channel or 850 °C in the average channel: these values are expected to limit the fuel temperature to levels which prevent excessive release of fission gases and high thermal stresses in the cooler (item 12 in Figure 1) and in the stainless steel cross ducts connecting the reactor vessel and the cooler (items from 6 to 11 in Figure 1) (Pagani et al., 2005).

Indicating by  $\mathbf{x}$  the vector of the 9 uncertain system parameters of Table 1 (Section 4.1) and by  $T_{out,core}^{hot}(\mathbf{x})$  and  $T_{out,core}^{avg}(\mathbf{x})$  the coolant outlet temperatures in the hot and average channels, respectively, the failure event  $F$  can be written as follows:

$$F = \{\mathbf{x} : T_{out,core}^{hot}(\mathbf{x}) > 1200\} \cup \{\mathbf{x} : T_{out,core}^{avg}(\mathbf{x}) > 850\}. \quad (4)$$

Notice that, in the notation of the preceding Section 3,  $T_{out,core}^{hot}(\mathbf{x}) = y_1(\mathbf{x})$  and  $T_{out,core}^{avg}(\mathbf{x}) = y_2(\mathbf{x})$  are the two target outputs of the T-H model.

The failure event  $F$  in (4) can be condensed into a single performance indicator  $Y(\mathbf{x})$  (Section 2) as

$$Y(\mathbf{x}) = \max \left\{ \frac{T_{out,core}^{hot}(\mathbf{x})}{1200}, \frac{T_{out,core}^{avg}(\mathbf{x})}{850} \right\} = \max \left\{ \frac{y_1(\mathbf{x})}{1200}, \frac{y_2(\mathbf{x})}{850} \right\} \quad (5)$$

so that the failure event  $F$  becomes specified as:

$$F = \{\mathbf{x} : Y(\mathbf{x}) > 1\}. \quad (6)$$

In the notation of Section 2, the failure threshold  $\alpha_Y$  is then equal to one.

## 5 Results of the application of bootstrapped ANNs for the functional failure analysis of the T-H passive system of Section 4

In this Section, the results of the application of bootstrapped Artificial Neural Networks (ANNs) for the quantitative functional failure analysis of the 600-MW GFR passive decay heat removal system in Figure 1 are illustrated. First, few details about the construction of the ANN regression model are given in Section 5.1; then, this is used to estimate the probability of functional failure of the system (Section 5.2); finally, the sensitivity of the hot-channel coolant outlet temperature to the uncertain input parameters is studied by computing first-order Sobol indices (Section 5.3). Notice that at each estimation step, the model uncertainties associated to the above mentioned quantities are also estimated by bootstrapping the ANN regression models (see Section 3 and the Appendix).

### 5.1 Building and testing the ANN regression model

The ANN model used in this work is built using a set  $D_{train} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{train}\}$  of input/output data examples of size  $N_{train} = 100$ ; this is done to test the capability of the ANN regression model to reproduce the outputs of the nonlinear T-H model code on the basis of a relatively *small* number of runs from the original T-H code. A Latin Hypercube Sample (LHS) of the inputs is drawn to give the vectors  $\mathbf{x}_p = \{x_{1,p}, x_{2,p}, \dots, x_{j,p}, \dots, x_{n_i,p}\}$ ,  $p = 1, 2, \dots, N_{train}$  (Zhang and Foschi, 2004). Then, the original T-H model is evaluated on the input vectors  $\mathbf{x}_p$ ,  $p = 1, 2, \dots, N_{train}$ , to obtain the corresponding output vectors  $\mathbf{y}_p = \boldsymbol{\mu}_y(\mathbf{x}_p) = \{y_{1,p}, y_{2,p}, \dots, y_{l,p}, \dots, y_{n_o,p}\}$ ,  $p = 1, 2, \dots, N_{train}$ , and build the data set  $D_{train} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{train}\}$ . Finally, the adjustable parameters  $\mathbf{w}^*$  of the ANN regression model are calibrated to fit the generated data: in particular, the common error back-propagation algorithm is implemented to *train* the ANN (Rumelhart et al., 1986).

In the present case study, the number  $n_i$  of inputs to the ANN regression model is equal to 9 (i.e., the number of uncertain inputs in Table 1 of Section 4.1), whereas the number  $n_o$  of outputs is equal to 2 (i.e., the number of system variables of interest, the hot- and average-channel coolant outlet temperatures, as reported in Section 4.2). The number of nodes  $n_h$  in the hidden layer has been set equal to 4 by trial and error.

A *validation* data set  $D_{val} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{val} = 20\}$  (different from the training set  $D_{train}$ ) is used to monitor the accuracy of the ANN model during the training procedure: in practice, the RMSE (2) is computed on  $D_{val}$  at different phases of the training procedure. At the beginning, the RMSE computed on the validation set  $D_{val}$  typically decreases together with the RMSE computed on the training set  $D_{train}$ ; then, when the ANN regression model starts overfitting the data, the RMSE calculated on the validation set  $D_{val}$  starts increasing: this is the time to stop the training algorithm.

For a realistic measure of the ANN model accuracy, the widely adopted coefficient of determination  $R^2$  and the RMSE have been computed for each output  $y_l$ ,  $l = 1, 2, \dots, n_o$ , on a new data set  $D_{test} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{test}\}$  also of size  $N_{test} = 20$ , not used during training (Marrel et al., 2009). Table 2 reports the values of the coefficient of determination  $R^2$  and of the RMSE associated to the estimates of the hot- and average- channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively, computed on the test set  $D_{test}$  of size  $N_{test} = 20$  by the ANN model with  $n_h = 4$  hidden neurons, built on a data set  $D_{train}$  of size  $N_{train} = 100$ ; the number of adjustable parameters (ANN weights)  $\mathbf{w}^*$  in the ANN regression model is also reported.

### Table 2 here

The large values of the coefficient of determination  $R^2$ , i.e., 0.9897 and 0.9866, and the small values of 12 °C and 6.3 °C for the RMSEs produced by the ANN for the hot- and average-channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively, lead us to assert that the accuracy of the ANN model can be considered satisfactory for the needs of estimating the functional failure probability of the present T-H passive system.

Finally, in order to demonstrate that the trial-and-error selected ANN architecture with  $n_h = 4$  hidden neurons is suitable for the present application, Table 2 contains also the values of the coefficient of determination  $R^2$  and of the RMSE associated to the estimates of the hot- and average- channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively, obtained on the test set  $D_{test}$  by two additional ANN topologies: in particular, ANN regression models with  $n_h = 3$  and 5 hidden neurons are considered. It can be seen that the values of the coefficient of determination  $R^2$  obtained by the ANN architecture with  $n_h = 3$  hidden neurons are 0.9821 and 0.9763 (i.e., lower than those produced by the ANN with  $n_h = 4$  hidden neurons), while the values of the RMSEs are 16.0 °C and 8.5 °C (i.e., larger than those produced by the ANN with  $n_h = 4$  hidden neurons) for the

hot- and average-channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively. The values of the coefficient of determination  $R^2$  obtained by the ANN architecture with  $n_h = 5$  hidden neurons are 0.9891 and 0.9860 (again, lower than those produced by the ANN with  $n_h = 4$  hidden neurons), while the values of the RMSEs are 13.4 °C and 7.6 °C (again, larger than those produced by the ANN with  $n_h = 4$  hidden neurons) for the hot- and average-channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively.

## 5.2 Functional failure probability estimation

In this Section, the bootstrapped ANNs are used for estimating the functional failure probability, and associated confidence interval of the 600-MW GFR passive decay heat removal system of Figure 1. The previous system configuration with  $N_{loops} = 3$  is analyzed.

Table 3 reports the value of the Bootstrap Bias Corrected (BBC) estimate  $\hat{P}(F)_{BBC}$  of the functional failure probability (per demand)  $P(F)$  obtained with  $N_T = 500000$  estimations from  $B = 1000$  bootstrapped ANNs built on  $N_{train} = 100$  data examples; the corresponding Bootstrap Bias Corrected (BBC)-95% Confidence Interval (CI) is also reported. A “true” value of the functional failure probability  $P(F)$  is also reported in Table 3 for reference (i.e.,  $P(F) = 3.34 \cdot 10^{-4}$ ); this has been obtained with a very large number  $N_T$  (i.e.,  $N_T = 500000$ ) of simulations of the original T-H code, which actually runs fast enough to allow repetitive calculations (one code run lasts on average 3 seconds on a Pentium 4 CPU 3.00GHz): the computational time required by this reference analysis is thus  $500000 \cdot 3 \text{ s} = 1500000 \text{ s} \approx 417 \text{ h}$ .

### Table 3 here

It can be seen that bootstrapped ANNs are quite *reliable* because the value of the BBC point estimate  $\hat{P}(F)_{BBC}$  (i.e.,  $3.59 \cdot 10^{-4}$ ) is quite close to the “true” value (i.e.,  $3.34 \cdot 10^{-4}$ ) of the functional failure probability  $P(F)$ . In spite of the small value of the failure probability (i.e.,  $P(F) \sim 10^{-4}$ ) this is done by resorting to quite a *low* number of runs of the T-H code (i.e., only  $N_{train} + N_{val} + N_{test} = 100 + 20 + 20 = 140$  input/output examples for training, validating and testing the bootstrapped ANN model).

In addition, the BBC-95% CI produced by the bootstrapped ANNs provides a measure of the (model) uncertainty associated to the ANN point estimate  $\hat{P}(F)_{BBC}$ : this information is particularly interesting when very few data are used to build the bootstrapped ANN models and, consequently,

the confidence of the analyst in the obtained BBC point estimate  $\hat{P}(F)_{BBC}$  is poor. In this respect, the upper bound of the BBC-95% CI (in this case,  $4.12 \cdot 10^{-4}$ ) can be used to provide a slightly *conservative* estimate of the corresponding functional failure probability  $P(F)$ . Note that the BBC-95% CI includes the “true” value of  $P(F)$  (i.e.,  $3.34 \cdot 10^{-4}$ ) and it is quite *narrow* around it: this confirms the *robustness* of the estimates produced by the trained bootstrapped ANN regression models in the present application.

Finally, a computational time of about 2 hours is associated to the calculation of the BBC point estimate  $\hat{P}(F)_{BBC}$  for  $P(F)$  and the corresponding BBC-95% CI (Table 3); this value includes the time required for i) generating the  $N_{train} + N_{val} + N_{test} = 100 + 20 + 20 = 140$  input/output examples, by running the T-H code (i.e., on average  $140 \cdot 3 \text{ s} = 420 \text{ s} = 7 \text{ m} \approx 0.12 \text{ h}$ ), ii) training the bootstrapped ensemble of  $B = 1000$  ANN models by means of an error back-propagation algorithm (i.e., on average 2 h) and iii) performing  $N_T = 500000$  evaluations of *each* of the  $B = 1000$  bootstrapped ANN models (i.e., on average 6 minutes = 0.1 h). The overall CPU time required by the use of bootstrapped ANNs (i.e., on average 2.22 h) is about 180 times lower than that required by the use of the original T-H model code (i.e., on average 417 h).

### 5.3 Global sensitivity analysis based on first-order Sobol indices

In the functional failure analysis of a T-H passive system, sensitivity analysis is a useful tool for identifying the uncertain parameters (i.e., the uncertain inputs to the T-H code) that contribute most to the variability of the model outputs (i.e., the coolant outlet temperatures). In the following, first-order Sobol sensitivity indices are computed only for the hot-channel coolant outlet temperature  $T_{out,core}^{hot}$  by way of example (Sobol, 1993).

By definition, the first-order Sobol sensitivity index  $S_j^l, j = 1, 2, \dots, n_i, l = 1, 2, \dots, n_o$ , quantifies the proportion of the variance of the output  $y_l, l = 1, 2, \dots, n_o$ , that can be attributed to the variance of the uncertain input variable  $x_j$  *alone*, i.e., *without* taking into account *interactions* with other input variables. It is formally defined as

$$S_j^l = \frac{V_{x_j} \left[ E_{x_{-j}} (y_l | x_j) \right]}{V[y_l]}, j = 1, 2, \dots, n_i, l = 1, 2, \dots, n_o, \quad (7)$$

where  $V[y_l]$  is the variance of the model output  $y_l, l = 1, 2, \dots, n_o$ , obtained when all model input parameters  $x_j, j = 1, 2, \dots, n_i$ , are sampled over their variation range,  $\mathbf{x}_{-j} = \{x_1, x_2, \dots, x_{j-1}, x_{j+1}, \dots, x_{n_i}\}$  is a vector containing all the uncertain input variables except  $x_j$  and  $V_{x_j} \left[ E_{x_{-j}} (y_l | x_j) \right]$  is the *expected variance* of  $y_l$  obtained when the input parameter of interest  $x_j$  is fixed and all the



remaining input variables are sampled over their variation ranges; a thorough description of this sensitivity measure goes far beyond the scope of this work: further details can be found in (Saltelli, 2002a).

As pointed out in (Saltelli, 2002b), the sensitivity index (7) has the advantage of being *global* because the effect of the *entire* distribution of the parameter whose uncertainty importance is evaluated, is considered; moreover, this sensitivity index is also “model free” because its computation is independent from assumptions about the model form, such as linearity, additivity and so on. The drawback of this approach relies in the computational burden associated to its calculation. Actually, if  $N$  samples (i.e., T-H model code evaluations) are used to calculate both the expected value  $E_{x_j}(y_l | x_j)$  and the variance  $V_{x_j}[E_{x_j}(y_l | x_j)]$  in (7) by crude Monte Carlo sampling, then the total number  $N_T$  of samples required to compute (7) is  $N_T = N^2$  for *each* output  $y_l$  of interest,  $l = 1, 2, \dots, n_o$  (Saltelli et al., 2008): for example, if  $N = 10^2 \div 10^3$ , then  $N_T = 10^4 \div 10^6$ , rendering impracticable the associated analysis. The total number  $N_T$  of samples can be reduced to  $N(n_i + 2)$  by resorting to an efficient algorithm proposed by (Saltelli, 2002b): in this case, if  $N = 10^2 \div 10^3$  and  $n_i = 9$  (like in the present problem), then  $N_T = 1.1 \cdot 10^3 \div 1.1 \cdot 10^4$ , leading to a reduction of one or two orders of magnitude in the number of required T-H code simulations; details about these algorithms can be found in the cited references.

In this work, the hot channel coolant outlet temperature  $T_{out,core}^{hot} = y_1$  is chosen as output of interest for the analysis and the algorithm proposed by (Saltelli, 2002b) has been implemented with  $N = 10000$  and  $n_i = 9$  (i.e.,  $N_T = 110000$ ) to obtain the “true” (i.e., reference) values of the first-order Sobol sensitivity indices  $S_j^1$  for the input variables  $x_j, j = 1, 2, \dots, 9$ . The reference ranking of the uncertain input variables obtained with  $N_T = 110000$  runs of the original T-H model code is reported in the left column of Table 4 together with the corresponding values of the first-order Sobol sensitivity indices  $S_j^1, j = 1, 2, \dots, 9$  (in parentheses) The Table also shows the ranking of the uncertain input parameters  $x_j, j = 1, 2, \dots, 9$ , based on the BBC point estimates  $\hat{S}_{j,BBC}^1, j = 1, 2, \dots, 9$ , obtained with  $N_T = 110000$  estimations from  $B = 1000$  bootstrapped ANN models built on  $N_{train} = 100$  input/output data examples (right column); finally, the computational time (in hours) associated to both analyses (i.e., on average 92 h and 2.12 h, respectively) is also reported, on a Pentium 4 CPU 3.00GHz.

**Table 4 here**

It can be seen that the ranking provided by bootstrapped ANNs is exactly the same as the reference one (i.e., the one obtained by running the original T-H model code a large number  $N_T$  of times): this results confirms the good estimation accuracy of the trained bootstrapped ANNs and the possibility to use this regression model for sensitivity analysis in T-H passive system functional failure assessment. It is interesting to note that the variances of the probability distributions of the first five uncertain input parameters/variables in the ranking (i.e.,  $x_2$ ,  $x_8$ ,  $x_5$ ,  $x_3$  and  $x_1$ ) accounts for about the 96% of the total variance of the probability distribution of the hot channel coolant outlet temperature  $T_{out,core}^{hot} = y_1$ : actually,  $S_2^1 + S_8^1 + S_5^1 + S_3^1 + S_1^1 = 0.9645$ . This outcome provides two important insights. On the one side, the analyst is able to identify those parameters/variables whose epistemic uncertainty plays a major role in determining the functional failure of the T-H passive system: consequently, his/her research efforts can be focused on increasing the state-of-knowledge only on these important parameters/variables and the related physical phenomena (for example, the collection of experimental data could lead to an improvement in the state-of-knowledge on the correlations used to model the heat transfer process in natural convection); on the other side, the analyst is allowed to identify those parameters/variables (in this case,  $x_6$ ,  $x_9$ ,  $x_4$  and  $x_7$ ) that are not important so that they may be excluded from the analysis, thereby simplifying the T-H model. In addition, it is worth recalling that in the present case these insights are obtained at the expense of only  $N_{train} + N_{val} + N_{test} = 100 + 20 + 20 = 140$  runs of the original T-H model code instead of the many thousands that are required by the crude and direct application of the algorithm in (Saltelli, 2002b) to the original, long-running T-H model code.

Finally, Table 5 reports the Bootstrap Bias Corrected (BBC)-95% Confidence Intervals (CIs) associated to the BBC point estimates  $\hat{S}_{j,BBC}^1$  of Table 4,  $j = 1, 2, \dots, 9$ . The information conveyed by these intervals is important when few data are used to train the bootstrapped ANNs and the consequent confidence of the analyst on the Sobol index point estimates  $\hat{S}_{j,BBC}^1$  is poor. In this respect, it is interesting to note that the *relative* width of the CIs of Sobol indices very close to one (i.e., those associated to very important input variables) is much lower with respect to those very close to zero (i.e., those associated to non important input variables): for example, the relative width of the BBC-95% CI of the first variable in the ranking, i.e.  $x_2$  (pressure), is  $(0.8324 - 0.7949)/0.8105 = 0.0463$ , whereas that of the fourth variable in the ranking, i.e.  $x_3$  (cooler wall temperature), is  $(0.0479 - 0.0352)/0.0303 = 0.4191$ . As expected, the *robustness* of the estimates of Sobol indices very close to zero is much lower than those of Sobol indices very close to one.

**Table 5 here**

## 6 Conclusions

In this paper, Artificial Neural Networks (ANNs) have been considered for performing a fast and efficient functional failure analysis of a T-H passive system. A case study involving the natural convection cooling in a Gas-cooled Fast Reactor (GFR) after a Loss of Coolant Accident (LOCA) has been taken as reference. For simplicity, the representation of the system behavior has been limited to a steady-state model.

An ANN model has been constructed on the basis of a limited-size set of data which represent examples of the nonlinear relationships between 9 uncertain inputs and 2 relevant outputs of the T-H model code (i.e., the hot- and average-channel coolant outlet temperatures). Once built, such model has been used as fast-running surrogate of the original, long-running T-H model code to perform the functional failure analysis of the T-H passive system. First, a functional failure probability as small as  $10^{-4}$  has been estimated; then, the sensitivity of the passive system performance to the uncertain system input parameters has been studied by calculating first-order Sobol sensitivity indices. In both analyses, the results have demonstrated that although the ANN regression model has been built on few (i.e., 100) data examples, the point estimates provided are *reliable*, because they are very near to the true values of the quantities of interest.

Moreover, a bootstrap of ANN regression models has been considered to produce confidence intervals for the estimates of the above mentioned safety quantities: this (model) uncertainty quantification is of paramount importance in safety critical applications, in particular when few data examples are used to build the surrogate models and, consequently, the confidence of the analyst in the obtained estimates is poor. With respect to that, the bootstrapped ANNs have been shown to be quite *robust* because the produced confidence intervals are satisfactorily *narrow* around the true values of the quantities of interest.

The results obtained show that the applied procedure is effective in reducing the computational burden associated to the functional failure analyses of T-H passive systems, while quantifying the uncertainty in the results. Although the T-H model used in this work to describe the behaviour of the natural circulation-based T-H passive system is a steady-state (thus, simplified) model, it is expected that even more significant benefits be gained with respect to more detailed thermal-hydraulic models (e.g., RELAP5-3D), provided that the number of code runs to train and validate the bootstrapped ANN regression model is small, as in the proposed procedure.

A final remark is also in order with respect to the possibility of using bootstrapped ANNs in the analysis of a *complete* accident sequence involving a T-H passive system instead of *only one phase* of the sequence (as it is done in the present work). In this view, two issues must be taken into account: i) the behavior of a T-H passive system is obviously dependent on the boundary conditions

of operation which depend on their turn on the particular phase of the accident sequence considered and on the “history” of the accident sequence itself: thus, possibly *different ANN regression models* should be built for use in *different phases* of the accident scenarios considered; ii) the (model) uncertainty associated to the estimates of the ANN regression model have to be propagated through the accident sequence: to this aim, the creation of bootstrap-based *empirical probability distributions* for the physical quantities of interest (e.g., the coolant temperature, the passive system failure probability and so on) offers a possible way to tackle this problem. Both issues i) and ii) above will be subject of future researches and applications.

## References

- Apostolakis, G. E., 1990. The concept of probability in safety assessment of technological systems. *Science*, 250, 1359.
- Arul, A.J., Iyer, N.K., Velusamy, K., 2009. Adjoint operator approach to functional reliability analysis of passive fluid dynamical systems. *Reliability Engineering and System Safety*, 94, pp. 1917-1926.
- Au, S. K. and Beck, J. L., 2001. Estimation of small failure probabilities in high dimensions by subset simulation. *Probabilistic Engineering Mechanics*, 16(4), pp. 263-277.
- Au, S. K. and Beck, J. L., 2003. Subset Simulation and its application to seismic risk based on dynamic analysis. *Journal of Engineering Mechanics*, 129(8), pp. 1-17.
- Bassi, C., Marquès, M., 2008. Reliability assessment of 2400 MWth gas-cooled fast reactor natural circulation decay heat removal in pressurized situations. *Science and Technology of Nuclear Installations*, Special Issue “Natural Circulation in Nuclear Reactor Systems”, Hindawi Publishing Corporation, Paper 87376.
- Baxt, W. G. and White, H., 1995. Bootstrapping confidence intervals for clinic input variable effects in a network trained to identify the presence of acute myocardial infarction. *Neural Computation*, 7, pp. 624-638.
- Bishop, C. M., 1995. *Neural Networks for pattern recognition*. Oxford University Press.
- Bucher, C., Most, T., 2008. A comparison of approximate response function in structural reliability analysis. *Probabilistic Engineering Mechanics*, vol. 23, pp. 154-163.
- Burgazzi, L., 2003. Reliability evaluation of passive systems through functional reliability assessment. *Nuclear Technology*, 144, 145.
- Burgazzi, L., 2007. State of the art in reliability of thermal-hydraulic passive systems. *Reliability Engineering and System Safety*, 92(5), pp. 671-675.

- Cacuci, D. G., Ionescu-Bujor, M., 2004. A comparative review of sensitivity and uncertainty analysis of large scale systems – II: Statistical methods. *Nuclear Science and Engineering* (147), pp. 204-217.
- Cadini, F., Zio, E., Kopustinskas, V., Urbonas, R., 2008. An empirical model based bootstrapped neural networks for computing the maximum fuel cladding temperature in a RBMK-1500 nuclear reactor accident. *Nuclear Engineering and Design*, 238, pp. 2165-2172.
- Cardoso, J. B., De Almeida, J. R., Dias, J. M., Coelho, P. G., 2008. Structural reliability analysis using Monte Carlo simulation and neural networks. *Advances in Engineering Software*, 39, pp. 505-513.
- Cheng, J., Li, Q. S., Xiao, R. C., 2008. A new artificial neural network-based response surface method for structural reliability analysis. *Probabilistic Engineering Mechanics*, 23, pp. 51-63.
- Deng, J., 2006. Structural reliability analysis for implicit performance function using radial basis functions. *International Journal of Solids and Structures*, 43, pp. 3255-3291.
- Efron, B. and Tibshirani, R. J., 1993. An introduction to the bootstrap. *Monographs on statistics and applied probability* 57. Chapman and Hall, New York.
- Fong, C. J., Apostolakis, G. E., Langewish, D. R., Hejzlar, P., Todreas, N. E., Driscoll, M. J., 2009. Reliability analysis of a passive cooling system using a response surface with an application to the flexible conversion ratio reactor. *Nuclear Engineering and Design*, doi:10.1016/j.nucengdes.2009.07.008.
- Gavin, H. P., Yau, S. C., 2008. High-order limit state functions in the response surface method for structural reliability analysis. *Structural Safety*, 30, pp. 162-179.
- Gazut, S., Martinez, J. M., Dreyfus, G., Oussar, Y., 2008. Towards the optimal design of numerical experiments. *IEEE Transactions on Neural Networks*, 19(5), pp. 874-882.
- Guba, A., Makai, M., Pal, L., 2003. Statistical aspects of best estimate method-I. *Reliability Engineering and System Safety*, 80, pp. 217-232.
- Helton, J. C., 1998. Uncertainty and sensitivity analysis results obtained in the 1996 performance assessment for the waste isolation power plant, SAND98-0365, Sandia National Laboratories.
- Helton, J., 2004. Alternative representations of epistemic uncertainties. *Reliability Engineering and System Safety*, 85 (Special Issue).
- Helton J. C, Davis, F. J., 2003. Latin hypercube sampling and the propagation of uncertainty in analyses of complex systems. *Reliability Engineering and System Safety*, 81, pp. 23-69.
- Helton, J. C., Johnson J. D., Sallaberry C. J., Storlie C. B., 2006. Survey on sampling-based methods for uncertainty and sensitivity analysis. *Reliability Engineering and System Safety*, 91, pp. 1175-1209.

- Hurtado, J. E., 2007. Filtered importance sampling with support vector margin: a powerful method for structural reliability analysis. *Structural Safety*, 29, pp. 2-15.
- IAEA, 1991. Safety related terms for advanced nuclear plant. IAEA TECDOC-626.
- Jafari, J., D' Auria, F., Kazeminejad, H., Davilu, H., 2003. Reliability evaluation of a natural circulation system. *Nuclear Engineering and Design*, 224, 79-104.
- Koutsourelakis, P. S., Pradlwarter, H. J., Schueller, 2004. Reliability of structures in high dimensions, Part I: algorithms and application. *Probabilistic Engineering Mechanics* (19), pp. 409-417.
- Liel, A. B., Haselton, C. B., Deierlein, G. G., Baker, J. W., 2008. Incorporating modeling uncertainties in the assessment of seismic collapse risk of buildings. *Structural Safety*, doi: 10.1016/j.strusafe.2008.06.002.
- Mackay F. J., Apostolakis, G. E., Hejzlar, P., 2008. Incorporating reliability analysis into the design of passive cooling systems with an application to a gas-cooled reactor. *Nuclear Engineering and Design*, 238(1), pp. 217-228.
- Makai, M., Pal, L., 2006. Best estimate method and safety analysis II. *Reliability Engineering and System Safety*, 91, pp. 222-232.
- Marquès, M., Pignatell, J. F., Saignes, P., D' Auria, F., Burgazzi, L., Müller, C., Bolado-Lavin, R., Kirchsteiger, C., La Lumia, V., Ivanov, I., 2005. Methodology for the reliability evaluation of a passive system and its integration into a probabilistic safety assessment. *Nuclear Engineering and Design*, 235, 2612-2631.
- Marrel, A., Iooss, B., Laurent, B., Roustant, O., 2009. Calculations of Sobol indices for the Gaussian process metamodel. *Reliability Engineering and System Safety*, vol., 94, pp. 742-751.
- Mathews, T. S., Ramakrishnan, M., Parthasarathy, U., John Arul, A., Senthil Kumar, C., 2008. Functional reliability analysis of safety grade decay heat removal system of Indian 500 MWe PFBR. *Nuclear Engineering and Design*, 238(9), pp. 2369-2376.
- Mathews, T.S., Arul, A.J., Parthasarathy, U., Kumar, C.S., Ramakrishnan, M., Subbaiah, K.V., 2009. Integration of functional reliability analysis with hardware reliability: An application to safety grade decay heat removal system of Indian 500 MWe PFBR. *Annals of Nuclear Energy*, 36, pp. 481-492.
- NUREG-1150, 1990. Severe accident risk: an assessment for five US nuclear power plants, US Nuclear Regulatory Commission.
- Nutt, W. T., Wallis, G. B., 2004. Evaluations of nuclear safety from the outputs of computer codes in the presence of uncertainties. *Reliability Engineering and System Safety*, 83, pp. 57-77.

- Pagani, L., Apostolakis, G. E. and Hejzlar, P., 2005. The impact of uncertainties on the performance of passive systems. *Nuclear Technology*, 149, 129-140.
- Patalano, G., Apostolakis, G. E., Hejzlar, P., 2008. Risk-informed design changes in a passive decay heat removal system. *Nuclear Technology*, vol. 163, pp. 191-208.
- Pradlwarter, H. J., Pellissetti, M. F., Schenk, C. A., Schueller, G. I., Kreis, A., Fransen, S., Calvi, A., Klein, M., 2005. *Computer Methods in Applied Mechanics and Engineering*, 194, pp. 1597-1617.
- Prosek, A., Cepin, M., 2008. Success criteria time windows of operator actions using RELAP5/MOD3.3 within human reliability analysis. *Journal of Loss Prevention in the Process Industries*, 21(3), pp. 260-267.
- Rumelhart, D. E., Hinton, G. E., Williams, R. J., 1986. Learning internal representations by error back-propagation. In Rumelhart, D. E. and McClelland, J. L. (Eds.), *Parallel distributed processing: exploration in the microstructure of cognition* (vol. 1). Cambridge (MA): MIT Press.
- Saltelli, A., 2002a. Sensitivity analysis for importance assessment. *Risk Analysis*, 22(3), pp. 579-590.
- Saltelli, A., 2002b. Making best use of model evaluations to compute sensitivity indices. *Comput. Phys. Commun.*, vol. 145, pp. 280-297.
- Saltelli, A., Ratto, M., Andres, T., Campolongo, F., Cariboni, J., Gatelli, D., Saisana, M., Tarantola, S., 2008. *Global sensitivity analysis. The Primer*. John Wiley and Sons Ltd.
- Schueller, G. I., 2007. On the treatment of uncertainties in structural mechanics and analysis. *Computers and Structures*, 85, pp. 235-243.
- Secchi, P., Zio, E., Di Maio, F., 2008. Quantifying uncertainties in the estimation of safety parameters by using bootstrapped artificial neural networks. *Annals of Nuclear Energy*, 35, pp. 2338-2350.
- Sobol, I. M., 1993. Sensitivity analysis for nonlinear mathematical model. *Math. Modelling Comput. Exp.*, 1, pp. 407-414.
- Storlie, C. B., Swiler, L. P., Helton, J. C., Sallaberry, C. J., 2008. Implementation and evaluation of nonparameteric regression procedures for sensitivity analysis of computationally demanding models. SANDIA Report n. SAND2008-6570.
- USNRC, 2002. "An approach for using probabilistic risk assessment in risk-informed decisions on plant-specific changes to the licensing basis." NUREG-1.174 – Revision 1, US Nuclear Regulatory Commission, Washington, DC.

- Volkova, E., Iooss, B., Van Dorpe, F., 2008. Global sensitivity analysis for a numerical model of radionuclide migration from the RRC “Kurchatov Institute” redwaste disposal site. *Stoch Environ Res Assess*, 22: pp. 17-31.
- Wilks, S. S., 1942. Statistical prediction with special reference to the problem of tolerance limits. *Annals of Mathematical Statistics*, 13, pp. 400-409.
- Zhang, J. and Foschi, R. O., 2004. Performance-based design and seismic reliability analysis using designed experiments and neural networks. *Probabilistic Engineering Mechanics*, 19., pp. 259-267.
- Zio, E., 2006. A study of the bootstrap method for estimating the accuracy of artificial neural networks in predicting nuclear transient processes. *IEEE Transactions on Nuclear Science*, 53(3), pp.1460-1470.
- Zio, E., and Apostolakis, G. E., 1996. Two methods for the structured assessment of model uncertainty by experts in performance assessment in radioactive waste repositories. *Reliability Engineering and System Safety*, Vol. 54, No. 2, 225-241.
- Zio, E. and Pedroni, N., 2009a. Estimation of the functional failure probability of a thermal-hydraulic passive systems by means of Subset Simulation. *Nuclear Engineering and Design*, 239, pp. 580-599.
- Zio, E. and Pedroni, N., 2009b. Functional failure analysis of a thermal-hydraulic passive system by means of Line Sampling. *Reliability Engineering and System Safety*, 94(11), pp. 1764-1781.

## Appendix

### The bootstrap algorithm for bias-corrected point and confidence interval estimation in ANN empirical regression modeling

In what follows, the steps of the procedure for the evaluation of the so-called Bootstrap Bias Corrected (BBC) point estimate  $\hat{Q}_{BBC}$  of a generic quantity  $Q$  (e.g., a safety parameter) by an ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  and the calculation of an associated BBC Confidence Interval (CI) are reported in detail (Zio, 2006; Storlie et al., 2008):

1. Generate a set  $D_{train}$  of input/output data examples by sampling  $N_{train}$  independent input parameters values  $\mathbf{x}_p, p = 1, 2, \dots, N_{train}$ , and calculating the corresponding set of  $N_{train}$  output vectors  $\mathbf{y}_p = \boldsymbol{\mu}_y(\mathbf{x}_p)$  through the mechanistic T-H system code. Plain random sampling, Latin Hypercube Sampling or other more sophisticated experimental design methods can be adopted to select the input vectors  $\mathbf{x}_p, p = 1, 2, \dots, N_{train}$  (Gazut et al., 2008).



2. Build an ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  on the basis of the entire data set  $D_{train} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{train}\}$  (step 1. above) in order to obtain a fast-running surrogate of the T-H model code represented by the unknown nonlinear deterministic function  $\mu_y(\mathbf{x})$  in (1).
3. Use the ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  (step 2. above), in place of the T-H model code, to provide a point estimate  $\hat{Q}$  of the quantity  $Q$ , e.g., the functional failure probability of the T-H passive system or a sensitivity index.

In particular, draw a sample of  $N_T$  new input vectors  $\mathbf{x}_r$ ,  $r = 1, 2, \dots, N_T$ , from the corresponding epistemic probability distributions and feed the ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  with them; then, use the corresponding output vectors  $\mathbf{y}_r = f(\mathbf{x}_r, \mathbf{w}^*)$ ,  $r = 1, 2, \dots, N_T$ , to calculate the estimate  $\hat{Q}$  for  $Q$  (the algorithm for computing  $\hat{Q}$  is obviously dependent on the meaning of the quantity  $Q$ ). Since the ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  can be evaluated quickly, this step is computationally costless even if the number  $N_T$  of model evaluations is very high (e.g.,  $N_T = 10^5$  or  $10^6$ ).

4. Build an ensemble of  $B$  (with  $B = 500-1000$ ) ANN regression models  $\{f_b(\mathbf{x}, \mathbf{w}_b^*), b=1, 2, \dots, B\}$  by random sampling with replacement and use each of the bootstrapped ANN regression models  $f_b(\mathbf{x}, \mathbf{w}_b^*)$ ,  $b = 1, 2, \dots, B$ , to calculate an estimate  $\hat{Q}_b$ ,  $b = 1, 2, \dots, B$ , for the quantity  $Q$  of interest: by so doing, a bootstrap-based empirical probability distribution for the quantity  $Q$  is produced which is the basis for the construction of the corresponding confidence intervals. In particular, repeat the following steps for  $b = 1, 2, \dots, B$ :

- a. Generate a bootstrap data set  $D_{train,b} = \{(\mathbf{x}_{p,b}, \mathbf{y}_{p,b}), p=1, 2, \dots, N_{train}\}$ ,  $b = 1, 2, \dots, B$ , by performing random sampling with replacement from the original data set  $D_{train} = \{(\mathbf{x}_p, \mathbf{y}_p), p=1, 2, \dots, N_{train}\}$  of  $N_{train}$  input/output patterns (steps 1. and 2. above). The data set  $D_{train,b}$  is thus constituted by the same number  $N_{train}$  of input/output patterns drawn among those in  $D_{train}$  although, due to the sampling with replacement, some of the patterns in  $D_{train}$  will appear more than once in  $D_{train,b}$ , whereas some will not appear at all.
- b. Build an ANN regression model  $f_b(\mathbf{x}, \mathbf{w}_b^*)$ ,  $b = 1, 2, \dots, B$ , on the basis of the bootstrap data set  $D_{train,b} = \{(\mathbf{x}_{p,b}, \mathbf{y}_{p,b}), p=1, 2, \dots, N_{train}\}$  (step 3.a. above).
- c. Use the ANN regression model  $f_b(\mathbf{x}, \mathbf{w}_b^*)$  (step 4.b. above), in place of the original T-H code, to provide a point estimate  $\hat{Q}_b$  of the quantity of interest  $Q$ . It is important to

note that for a correct quantification of the confidence interval the estimate  $\hat{Q}_b$  must be based on the same input and output vectors  $\mathbf{x}_r$  and  $\mathbf{y}_r$ ,  $r = 1, 2, \dots, N_T$ , respectively, obtained in step 3. above.

5. Calculate the so-called Bootstrap Bias Corrected (BBC) point estimate  $\hat{Q}_{BBC}$  for  $Q$  as

$$\hat{Q}_{BBC} = 2\hat{Q} - \hat{Q}_{boot} \quad (1')$$

where  $\hat{Q}$  is the estimate obtained with the ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  trained with the original data set  $D_{train}$  (steps 2. and 3. above) and  $\hat{Q}_{boot}$  is the average of the  $B$  estimates  $\hat{Q}_b$  obtained with the  $B$  ANN regression models  $f_b(\mathbf{x}, \mathbf{w}_b^*)$ ,  $b = 1, 2, \dots, B$  (step 4.c. above), i.e.,

$$\hat{Q}_{boot} = \frac{1}{B} \sum_{b=1}^B \hat{Q}_b. \quad (2')$$

The BBC estimate  $\hat{Q}_{BBC}$  in (1') is taken as the point estimate for  $Q$ .

The explanation for expression (1') is as follows. It can be demonstrated that if there is a bias in the bootstrap average estimate  $\hat{Q}_{boot}$  in (2') compared to the estimate  $\hat{Q}$  obtained with the single ANN regression model  $f(\mathbf{x}, \mathbf{w}^*)$  (step 3. above), then the same bias exists in the single estimate  $\hat{Q}$  compared to the true value  $Q$  of the quantity of interest (Baxt and White, 1995). Thus, in order to obtain an appropriate, i.e. bias-corrected, estimate  $\hat{Q}_{BBC}$  for the quantity of interest  $Q$ , the estimate  $\hat{Q}$  must be adjusted by subtracting the corresponding bias ( $\hat{Q}_{boot} - \hat{Q}$ ): as a consequence, the final, bias-corrected estimate  $\hat{Q}_{BBC}$  is  $\hat{Q}_{BBC} = \hat{Q} - (\hat{Q}_{boot} - \hat{Q}) = 2\hat{Q} - \hat{Q}_{boot}$ .

6. Calculate the two-sided Bootstrap Bias Corrected (BBC)- $100 \cdot (1 - \alpha)\%$  Confidence Interval (CI) for the BBC point estimate in (1') by performing the following steps:
- Order the bootstrap estimates  $\hat{Q}_b$ ,  $b = 1, 2, \dots, B$ , (step 4.c. above) by increasing values, such that  $\hat{Q}_{(i)} = \hat{Q}_b$  for some  $b = 1, 2, \dots, B$ , and  $\hat{Q}_{(1)} < \hat{Q}_{(2)} < \dots < \hat{Q}_{(b)} < \dots < \hat{Q}_{(B)}$ .
  - Identify the  $100 \cdot \alpha/2^{\text{th}}$  and  $100 \cdot (1 - \alpha/2)^{\text{th}}$  quantiles of the bootstrapped empirical probability distribution of  $Q$  (step 4. above) as the  $[B \cdot \alpha/2]^{\text{th}}$  and  $[B(1 - \alpha/2)]^{\text{th}}$  elements  $\hat{Q}_{([B \cdot \alpha/2])}$  and  $\hat{Q}_{([B(1 - \alpha/2)])}$ , respectively, in the ordered list  $\hat{Q}_{(1)} < \hat{Q}_{(2)} < \dots < \hat{Q}_{(b)} < \dots < \hat{Q}_{(B)}$ ; notice that the symbol  $[\cdot]$  stands for "closest integer".
  - Calculate the two-sided BBC- $100 \cdot (1 - \alpha)\%$  CI for  $\hat{Q}_{BBC}$  as

$$\left[ \hat{Q}_{BBC} - \left( \hat{Q}_{boot} - \hat{Q}_{\lfloor B \cdot \alpha/2 \rfloor} \right), \hat{Q}_{BBC} + \left( \hat{Q}_{\lfloor B \cdot (1-\alpha/2) \rfloor} - \hat{Q}_{boot} \right) \right]. \quad (3')$$

An important advantage of the bootstrap method is that it provides confidence intervals for a given quantity  $Q$  without making any model assumptions (e.g., normality); a disadvantage is that the computational cost could be high when the set  $D_{train}$  and the number of adaptable parameters  $\mathbf{w}^*$  in the regression models are large.

**FIGURE**

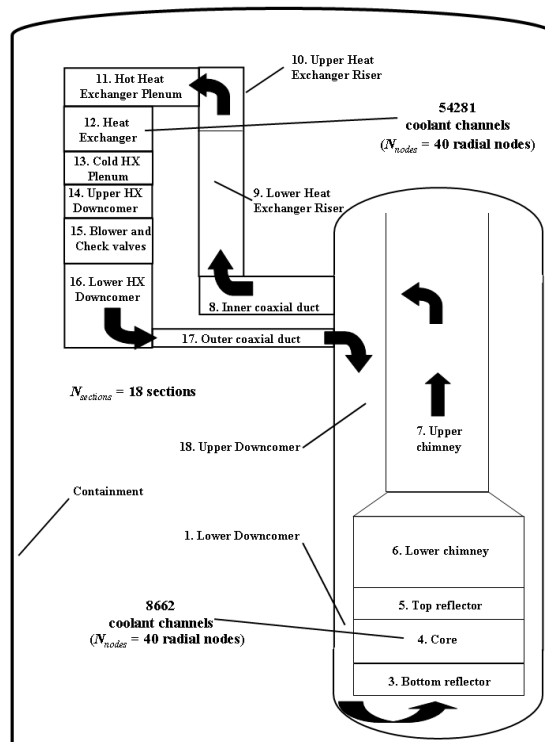


Figure 1. Schematic representation of one loop of the 600-MW GFR passive decay heat removal system (Pagani et al., 2005)

## TABLES

	Name	Mean, $\mu$	Standard deviation, $\sigma$ (% of $\mu$ )
<b>Parameter uncertainty</b>	Power (MW), $x_1$	18.7	1%
	Pressure (kPa), $x_2$	1650	7.5%
	Cooler wall temperature ( $^{\circ}\text{C}$ ), $x_3$	90	5%
<b>Model uncertainty (error factor, <math>\zeta</math>)</b>	Nusselt number in forced convection, $x_4$	1	5%
	Nusselt number in mixed convection, $x_5$	1	15%
	Nusselt number in free convection, $x_6$	1	7.5%
	Friction factor in forced convection, $x_7$	1	1%
	Friction factor in mixed convection, $x_8$	1	10%
	Friction factor in free convection, $x_9$	1	1.5%

*Table 1. Epistemic uncertainties considered for the 600-MW GFR passive decay heat removal system of Figure 1 (Pagani et al., 2005)*

<b>Artificial Neural Network (ANN)</b>							
<b>Optimal configuration selected: <math>n_i = 9, n_h = 4, n_o = 2</math></b>							
				$R^2$		RMSE [ $^{\circ}\text{C}$ ]	
$N_{train}$	$N_{val}$	$N_{test}$	Number of adjustable parameters $w^*$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$
100	20	20	50	0.9897	0.9866	12.0	6.3
<b>Configuration: <math>n_i = 9, n_h = 3, n_o = 2</math></b>							
				$R^2$		RMSE [ $^{\circ}\text{C}$ ]	
$N_{train}$	$N_{val}$	$N_{test}$	Number of adjustable parameters $w^*$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$
100	20	20	38	0.9821	0.9763	16.0	8.5
<b>Configuration: <math>n_i = 9, n_h = 5, n_o = 2</math></b>							
				$R^2$		RMSE [ $^{\circ}\text{C}$ ]	
$N_{train}$	$N_{val}$	$N_{test}$	Number of adjustable parameters $w^*$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$	$T_{out,core}^{hot}$	$T_{out,core}^{avg}$
100	20	20	62	0.9891	0.9860	13.4	7.6

*Table 2. Coefficient of determination  $R^2$  and RMSE associated to the ANN estimates of the hot- and average-channel coolant outlet temperatures  $T_{out,core}^{hot}$  and  $T_{out,core}^{avg}$ , respectively*

Failure probability (“True” value, $P(F) = 3.34 \cdot 10^{-4}$ ; CPU time $\approx 417$ h)					
Bootstrapped Artificial Neural Networks (ANNs)					
$N_{train}$	$N_{val}$	$N_{test}$	BBC point estimate, $\hat{P}(F)_{BBC}$	BBC-95% CI	CPU time
100	20	20	$3.59 \cdot 10^{-4}$	$[2.55 \cdot 10^{-4}, 4.12 \cdot 10^{-4}]$	$\approx 2.22$ h

Table 3. Bootstrap Bias Corrected (BBC) point estimate  $\hat{P}(F)_{BBC}$  and BBC-95% Confidence Interval (CI) of the functional failure probability  $P(F)$

Ranking of the uncertain input variables – Hot-channel coolant outlet temperature, $T_{out,core}^{hot} = y_1$	
Original T-H code, $N_T = 110000$ ( $S_j^1$ ); CPU time $\approx 92$ h	Bootstrapped ANNs, $N_{train} = 100$ ( $\hat{S}_{j,BBC}^1$ ); CPU time $\approx 2.12$ h
Pressure, $x_2$ (0.8105)	Pressure, $x_2$ (0.8098)
Friction mixed, $x_8$ (0.0594)	Friction mixed, $x_8$ (0.0605)
Nusselt mixed, $x_5$ (0.0583)	Nusselt mixed, $x_5$ (0.0591)
Cooler wall temperature, $x_3$ (0.0303)	Cooler wall temperature, $x_3$ (0.0368)
Power, $x_1$ ( $5.950 \cdot 10^{-3}$ )	Power, $x_1$ ( $5.199 \cdot 10^{-3}$ )
Nusselt free, $x_6$ ( $5.211 \cdot 10^{-4}$ )	Nusselt free, $x_6$ ( $6.338 \cdot 10^{-4}$ )
Friction free, $x_9$ ( $2.139 \cdot 10^{-4}$ )	Friction free, $x_9$ ( $1.676 \cdot 10^{-4}$ )
Nusselt forced, $x_4$ ( $4.214 \cdot 10^{-5}$ )	Nusselt forced, $x_4$ ( $6.430 \cdot 10^{-5}$ )
Friction forced, $x_7$ ( $1.533 \cdot 10^{-5}$ )	Friction forced, $x_7$ ( $1.634 \cdot 10^{-5}$ )

Table 4. Ranking of the uncertain input variables  $x_j$ ,  $j = 1, 2, \dots, 9$ , based on the first-order Sobol sensitivity indices  $S_j^1$  calculated for the hot channel coolant outlet temperature  $T_{out,core}^{hot} = y_1$

<b>First-order Sobol sensitivity indices <math>S_j^1</math> - Hot-channel coolant outlet temperature, <math>T_{out,core}^{hot} = y_1</math></b>		
<b>Bootstrapped Artificial Neural Networks (ANNs), <math>N_{train} = 100</math>; CPU time <math>\approx 2.12</math> h</b>		
<b>Variable, <math>x_j</math> (<math>S_j^1</math>)</b>	<b>BBC point estimate, <math>\hat{S}_{j,BBC}^1</math></b>	<b>BBC-95% CI</b>
Power, $x_1$ ( $5.950 \cdot 10^{-3}$ )	$5.199 \cdot 10^{-3}$	$[4.137 \cdot 10^{-4}, 8.563 \cdot 10^{-3}]$
Pressure, $x_2$ (0.8105)	0.8098	[0.7949, 0.8324]
Cooler wall temperature, $x_3$ (0.0303)	0.0368	[0.0352, 0.0479]
Nusselt forced, $x_4$ ( $4.214 \cdot 10^{-5}$ )	$6.430 \cdot 10^{-5}$	$[0, 9.523 \cdot 10^{-5}]$
Nusselt mixed, $x_5$ (0.0583)	0.0591	[0.0491, 0.0649]
Nusselt free, $x_6$ ( $5.211 \cdot 10^{-4}$ )	$6.338 \cdot 10^{-4}$	$[0, 8.413 \cdot 10^{-4}]$
Friction forced, $x_7$ ( $1.533 \cdot 10^{-5}$ )	$1.634 \cdot 10^{-5}$	$[0, 4.393 \cdot 10^{-5}]$
Friction mixed, $x_8$ (0.0594)	0.0605	[0.0536, 0.0711]
Friction free, $x_9$ ( $2.139 \cdot 10^{-4}$ )	$1.676 \cdot 10^{-4}$	$[0, 3.231 \cdot 10^{-4}]$

*Table 5. Bootstrap Bias Corrected (BBC) point estimates  $\hat{S}_{j,BBC}^1$ ,  $j = 1, 2, \dots, 9$ , and BBC-95% Confidence Intervals (CIs) of the first-order Sobol sensitivity indices  $S_j^1$ ,  $j = 1, 2, \dots, 9$ , calculated for the hot channel coolant outlet temperature  $T_{out,core}^{hot} = y_1$*