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► **To cite this version:**

Jakob Hoydis, Merouane Debbah, Mari Kobayashi. Asymptotic Moments for Interference Mitigation in Correlated Fading Channels. 2011 IEEE International Symposium on Information Theory Proceedings, Jul 2011, St. Petersburg, Russia. pp.2796 - 2800, 10.1109/ISIT.2011.6034083 . hal-00648020

HAL Id: hal-00648020

<https://hal-supelec.archives-ouvertes.fr/hal-00648020>

Submitted on 4 Dec 2011

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Asymptotic Moments for Interference Mitigation in Correlated Fading Channels

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Abstract—We consider a certain class of large random matrices, composed of independent column vectors with zero mean and different covariance matrices, and derive asymptotically tight deterministic approximations of their moments. This random matrix model arises in several wireless communication systems of recent interest, such as distributed antenna systems or large antenna arrays. Computing the linear minimum mean square error (LMMSE) detector in such systems requires the inversion of a large covariance matrix which becomes prohibitively complex as the number of antennas and users grows. We apply the derived moment results to the design of a low-complexity polynomial expansion detector which approximates the matrix inverse by a matrix polynomial and study its asymptotic performance. Simulation results corroborate the analysis and evaluate the performance for finite system dimensions.

I. INTRODUCTION

Distributed antenna systems and large antenna arrays have recently attained significant research interest [1], [2]. Both are considered as promising solutions to counter intercell interference and to increase the spectral efficiency of current cellular networks. Since these techniques rely in essence on a significant increase of the number of *coordinated antennas*, the computational complexity of the joint precoding/detection of the transmitted/received signals grows. This calls for low-complexity solutions. In this paper, we address this need by assessing the performance of a polynomial expansion detector [3] adapted to the following general channel model.

Consider a discrete-time $N \times K$ multiple-input multiple-output (MIMO) channel with output vector $\mathbf{y} \in \mathbb{C}^N$:

$$\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n} \quad (1)$$

where $\mathbf{x} = [x_1, \dots, x_K]^T$ is the complex channel input vector satisfying $\mathbb{E}[\mathbf{x}\mathbf{x}^H] = \mathbf{I}_K$, $\mathbf{H} = [\mathbf{h}_1 \dots \mathbf{h}_K] \in \mathbb{C}^{N \times K}$ is the random channel matrix and $\mathbf{n} \sim \mathcal{CN}(\mathbf{0}, \sigma^2 \mathbf{I}_N)$ is a vector of additive noise. The j th column $\mathbf{h}_j \in \mathbb{C}^N$ of \mathbf{H} is modeled as

$$\mathbf{h}_j = \frac{1}{\sqrt{K}} \mathbf{R}_j \mathbf{w}_j, \quad j = 1, \dots, K \quad (2)$$

where $\mathbf{R}_j \in \mathbb{C}^{N \times N}$ is a deterministic matrix and the elements of $\mathbf{w}_j \in \mathbb{C}^N$ are independent and identically distributed (i.i.d.) random variables with zero mean, unit variance and finite eighth moment. This channel model captures different types of wireless communication systems and generalizes several well-known channel models as discussed below:

Distributed Antenna Systems: Let $\mathbf{R}_j = \text{diag}(r_{1j}, \dots, r_{Nj})$ with elements $r_{ij} = \sqrt{p_j}/d_{ij}^{\beta/2}$, where d_{ij} is the (normalized) distance between transmitter j and receive antenna i , β is the path loss exponent and p_j is the transmit power of transmitter j . This model is suitable for distributed antenna systems [1] where each transmitter sees a different path loss to each of the receive antennas since d_{1j}, \dots, d_{Nj} are different.

Large-scale MIMO: Assume a receiver equipped with a very large antenna array ($N \gg 1$) as in [2]. Unless the antenna spacing is sufficiently large, it is likely that the received signals at different receive antennas are correlated. Our model allows to assign a different correlation matrix \mathbf{R}_j to each transmitter.

MIMO Multiple Access Channel (MAC): Consider a MIMO MAC from M transmitters equipped with $K_m, m = 1, \dots, M$, antennas to a receiver with N antennas. Each point-to-point link has a different transmit and receive correlation matrix [4]:

$$\mathbf{y} = \sum_{m=1}^M \Phi_{\mathbf{R},m}^{\frac{1}{2}} \mathbf{W}_m \Phi_{\mathbf{T},m}^{\frac{1}{2}} \mathbf{x}_m + \mathbf{n}$$

where $\Phi_{\mathbf{R},1}, \dots, \Phi_{\mathbf{R},M} \in \mathbb{C}^{N \times N}$ are deterministic correlation matrices, $\Phi_{\mathbf{T},1} \in \mathbb{C}^{K_1 \times K_1}, \dots, \Phi_{\mathbf{T},M} \in \mathbb{C}^{K_M \times K_M}$ are nonnegative diagonal matrices, $\mathbf{W}_1 \in \mathbb{C}^{N \times K_1}, \dots, \mathbf{W}_M \in \mathbb{C}^{N \times K_M}$ are random channel matrices with i.i.d. entries with zero mean and variance $1/K$, and $\mathbf{x}_1 \in \mathbb{C}^{K_1}, \dots, \mathbf{x}_M \in \mathbb{C}^{K_M}$ are the transmit vectors. Let $\sum_{m=1}^M K_m = K$. Setting $\mathbf{R}_j = \Phi_{\mathbf{R},m}^{1/2} [\Phi_{\mathbf{T},m}^{1/2}]_{ii}$ for $j \in \{1 + \sum_{l=1}^{m-1} K_l, \dots, \sum_{l=1}^m K_l\}$ and $i = j - \sum_{l=1}^{m-1} K_l$, we fall back to the model in (2).

In the sequel, we will study the asymptotic behavior of the moments μ_n of the matrix $\mathbf{B} \triangleq \mathbf{H}\mathbf{H}^H$, defined as

$$\mu_n \triangleq \frac{1}{N} \text{tr} \mathbf{B}^n, \quad n = 0, 1, 2, \dots \quad (3)$$

under the assumption that N and K grow infinitely large at the same speed. In particular, we will derive deterministic approximations $\bar{\mu}_n$ of μ_n , such that $\mu_n - \bar{\mu}_n \rightarrow 0$ almost surely, for $N, K \rightarrow \infty$. This result can be used, for example, to compute low-complexity approximations of the matrix inverse $(\mathbf{B} + \sigma^2 \mathbf{I}_N)^{-1}$. The computation of this matrix arises in many practical applications, such as for linear multiuser detectors and beamforming strategies. We will focus exemplarily on the linear minimum mean square error (LMMSE) detector.

The LMMSE estimate $\hat{\mathbf{x}}$ of \mathbf{x} , assuming perfect knowledge of \mathbf{H} at the receiver, is given as [5]

$$\hat{\mathbf{x}} = \mathbf{H}^H (\mathbf{B} + \sigma^2 \mathbf{I}_N)^{-1} \mathbf{y}. \quad (4)$$

The computational complexity of this estimate is of order $\mathcal{O}(r^2)$ [6], where $r = \min(N, K)$. A reduced complexity estimate can be obtained by approximating the matrix inverse in (4) by the following matrix polynomial [3]

$$(\mathbf{B} + \sigma^2 \mathbf{I}_N)^{-1} \approx \sum_{l=0}^{L-1} w_l \mathbf{B}^l \quad (5)$$

for some coefficients w_l , where the filter rank $L \leq r$ is chosen according to the allowable complexity. For a given transmitter k , the above *polynomial expansion detector* can be seen as a projection of \mathbf{y} on the L th Krylov subspace associated to the pair $(\mathbf{B}, \mathbf{h}_k)$, i.e., the subspace of \mathbb{C}^N spanned by the vectors $\{\mathbf{h}_k, \mathbf{B}\mathbf{h}_k, \dots, \mathbf{B}^{L-1}\mathbf{h}_k\}$, and a weighting of the joint projections by the coefficients w_l . Depending on L , the polynomial expansion detector achieves a performance between the matched filter ($L = 1$) and the LMMSE detector ($L = r$) [3] and allows, consequently, to trade-off performance for complexity. Moreover, (5) allows for an efficient multistage implementation [3], [7], [6], where each stage l consists of a matched filter \mathbf{H}^H and subsequent “re-spreading” by the matrix \mathbf{H} . In [8], it was shown that the signal-to-interference-plus-noise ratio (SINR) at the filter output converges in certain cases exponentially in the filter rank L to the SINR output of the LMMSE detector. Thus, L does not need to scale with the system size to achieve close to optimal performance [9].

The optimal weight vector $\mathbf{w} = [w_0 \dots w_{L-1}]^T$ can be chosen to minimize the mean square error of the estimated vector $\hat{\mathbf{x}}$, i.e.,

$$\mathbf{w} = \arg \min_{\mathbf{u}=[u_0, \dots, u_{L-1}]^T} \mathbb{E} \left[\left\| \mathbf{x} - \mathbf{H}^H \sum_{l=0}^{L-1} u_l \mathbf{B}^l \mathbf{y} \right\|_2^2 \right]. \quad (6)$$

The solution to this optimization problem is given as [3]

$$\mathbf{w} = \mathbf{\Phi}^{-1} \boldsymbol{\varphi} \quad (7)$$

where $\mathbf{\Phi} \in \mathbb{R}_+^{L \times L}$ and $\boldsymbol{\varphi} \in \mathbb{R}_+^L$ are defined as

$$\begin{aligned} [\mathbf{\Phi}]_{ij} &= \mu_{i+j} + \sigma^2 \mu_{i+j-1} \\ [\boldsymbol{\varphi}]_i &= \mu_i. \end{aligned} \quad (8)$$

The computation of the weight vector \mathbf{w} requires the calculation of the moments μ_1, \dots, μ_{2L} which is still computational expensive for large L . However, under the assumption that the dimensions of \mathbf{H} grow infinitely large, it was shown for several random matrix models (e.g. [7], [9], [10]) that the moments μ_n can be closely approximated by their asymptotic counterparts $\bar{\mu}_n$. These are independent of a particular realization of \mathbf{H} and can be calculated based on the statistical properties of the channel matrix. If these properties change on a much slower timescale than the fast-fading channel fluctuations, the weight vector \mathbf{w} can be precomputed using $\bar{\mu}_n$ instead of μ_n . Thus, the detector complexity depends only on the complexity of the projection on the Krylov subspace which is of order $\mathcal{O}(r)$ [6].

Multistage or reduced-rank multiuser detectors were mainly considered in the context of code-division multiple-access (CDMA) systems as low-complexity solutions to the joint detection of a large number of user terminals with long spreading sequences [3]. The asymptotic (universal) weight design was first studied in [7] for the equal transmit power case and then extended to more involved models, such as different transmit powers [9], [?], multi-path fading [10] and random unitary spreading sequences [11]. These results were then put on a common ground in [6] which compares different types of linear multistage detectors in terms of their complexity and asymptotic performance. Recently, also multistage detectors for asynchronous CDMA systems were considered in [12].

The asymptotic results in the above works are based on the almost sure (a.s.) convergence of the empirical spectral distribution (e.s.d.) of the matrix \mathbf{B} to a compactly supported limit distribution. This limit distribution is in general given implicitly by its Stieltjes transform which can be computed based on the statistical properties of the underlying random matrix model. The asymptotic moments are then obtained by writing the Stieltjes transform as a moment generating function [13, Theorem 2.3] and relying on combinatorial arguments [10] or free probability theory [11].

The technique used in this work is different in two aspects. First, we do not require the existence of a limiting eigenvalue distribution of the matrix \mathbf{B} . Instead, we provide for each pair (N, K) a deterministic approximation $\bar{\mu}_n$ of the moments μ_n which becomes arbitrarily tight as $N, K \rightarrow \infty$. Second, the moments are derived through iterated differentiation of the Stieltjes transform and can be computed by simple recursive equations. This is in contrast to [10] which requires an exhaustive search over complicated sets of indices. Hence, our results are more practical from an implementation perspective. Moreover, the asymptotic moments of the random matrix model (2) have not been considered in the literature before.

The paper is structured as follows: Section II contains definitions and related results. The asymptotic moments of \mathbf{B} are derived in Section III and the performance of the polynomial expansion receiver is studied in Section IV. Numerical results are provided in Section V. Section VI concludes the paper.

II. RELATED RESULTS

We need the following definitions and related results. Denote by “ \Rightarrow ” and “ $\xrightarrow{\text{a.s.}}$ ” weak and almost sure convergence.

Definition 1 (Empirical spectral distribution): Let $\mathbf{A} \in \mathbb{C}^{N \times N}$ be a Hermitian matrix with eigenvalues $\lambda_1, \dots, \lambda_N$. Denote $F^{\mathbf{A}}$ the e.s.d. of \mathbf{A} , defined as

$$F^{\mathbf{A}}(x) = \frac{1}{N} \sum_{i=1}^N \mathbb{1}(\lambda_i \leq x).$$

Definition 2 (Stieltjes transform): Let F be a real measurable function over \mathbb{R} with support $\text{Supp}(F)$. For $z \in \mathbb{C} \setminus \text{Supp}(F)$, the Stieltjes transform $m_F(z)$ of F is defined as

$$m_F(z) = \int_{-\infty}^{\infty} \frac{1}{\lambda - z} dF(\lambda).$$

Denote by \mathcal{S} the class of functions f analytic over $\mathbb{C} \setminus \mathbb{R}_+$, such that, for $z \in \mathbb{C}_+$, $f \in \mathbb{C}_+$, $zf \in \mathbb{C}_+$ and $\lim_{y \rightarrow \infty} -iyf(iy) < \infty$. Such functions are known to be Stieltjes transforms of finite measures supported by \mathbb{R}_+ [13, Theorem 2.2].

Theorem 1 ([14, Theorem 1]): Let $\mathbf{D} \in \mathbb{C}^{N \times N}$ be a Hermitian non-negative definite matrix and assume that \mathbf{D} and the matrices \mathbf{R}_j , $j = 1, \dots, K$, have uniformly bounded spectral norms (with respect to N). Let $N, K \rightarrow \infty$, such that $0 < \liminf \frac{K}{N} \leq \limsup \frac{K}{N} < \infty$. Then, for any $z \in \mathbb{C} \setminus \mathbb{R}_+$,

$$\frac{1}{N} \text{tr} \mathbf{D} (\mathbf{B} - z \mathbf{I}_N)^{-1} - \frac{1}{N} \text{tr} \mathbf{D} \mathbf{T}(z) \xrightarrow{\text{a.s.}} 0$$

where $\mathbf{T}(z) \in \mathbb{C}^{N \times N}$ is defined as

$$\mathbf{T}(z) \triangleq \left(\frac{1}{K} \sum_{j=1}^K \frac{\mathbf{R}_j \mathbf{R}_j^H}{1 + \delta_j(z)} - z \mathbf{I}_N \right)^{-1} \quad (9)$$

and the following set of K implicit equations

$$\delta_j(z) = \frac{1}{K} \text{tr} \mathbf{R}_j \mathbf{R}_j^H \mathbf{T}(z), \quad j = 1, \dots, K$$

admits a unique solution $(\delta_1(z), \dots, \delta_K(z)) \in \mathcal{S}^K$. Moreover, denote by F the distribution function whose Stieltjes transform is given by $m(z) = \frac{1}{N} \text{tr} \mathbf{T}(z)$. Then, almost surely,

$$F^{\mathbf{B}} - F \Rightarrow 0.$$

III. ASYMPTOTIC MOMENTS

In this section, we state our main results. Outlines of the proofs of Theorems 2 and 3 are provided in the appendix.

Theorem 2: Let F be the distribution function as defined in Theorem 1 and denote by $\bar{\mu}_0, \bar{\mu}_1, \dots$ the successive moments of F , defined as $\bar{\mu}_n \triangleq \int_0^\infty \lambda^n dF(\lambda)$. These moments can be calculated as

$$\bar{\mu}_n = \frac{(-1)^n}{n!} \frac{1}{N} \text{tr} \mathbf{T}_n$$

where \mathbf{T}_n is defined recursively by the following set of equations for $n \geq 0$:

$$\begin{aligned} \mathbf{T}_{n+1} &= \sum_{i=0}^n \sum_{j=0}^i \binom{n}{i} \binom{i}{j} \mathbf{T}_{n-i} \mathbf{Q}_{i-j+1} \mathbf{T}_j \\ \mathbf{Q}_{n+1} &= \frac{n+1}{K} \sum_{k=1}^K f_{k,n} \mathbf{R}_k \mathbf{R}_k^H \\ f_{k,n+1} &= \sum_{i=0}^n \sum_{j=0}^i \binom{n}{i} \binom{i}{j} (n-i+1) f_{k,j} f_{k,i-j} \delta_{k,n-i} \\ \delta_{k,n+1} &= \frac{1}{K} \text{tr} \mathbf{R}_k \mathbf{R}_k^H \mathbf{T}_{n+1} \end{aligned}$$

where $\mathbf{T}_0 = \mathbf{I}_N$, $f_{k,0} = -1$ and $\delta_{k,0} = \frac{1}{K} \text{tr} \mathbf{R}_k \mathbf{R}_k^H \forall k$.

Remark 3.1: While Theorem 2 allows to compute the moments $\bar{\mu}_n$ of F , it does not imply the a.s. convergence of μ_n and $\bar{\mu}_n$ in general. Theorem 3 provides some sufficient conditions for which this convergence holds.

Remark 3.2: Although difficult to show analytically, one can verify numerically that Theorem 2 coincides with [10, Theorem 1] for $\mathbf{R}_j = \text{diag}(r_{1j}, \dots, r_{Nj})$, $j = 1, \dots, K$.

If the matrices \mathbf{R}_j are drawn from a finite set of matrices, we get the following stronger result:

Theorem 3: For fixed $M > 0$, let $\mathcal{R} = \{\tilde{\mathbf{R}}_1, \dots, \tilde{\mathbf{R}}_M\}$ be a set of complex $N \times N$ matrices and let $\mathbf{D} \in \mathbb{C}^{N \times N}$ be a non-negative definite Hermitian matrix. Assume that \mathbf{D} and $\tilde{\mathbf{R}}_m$, $m = 1, \dots, M$, have uniformly bounded spectral norms (with respect to N). Let $\mathbf{R}_j \in \mathcal{R}$ for $j = 1, \dots, K$. Assume $N, K \rightarrow \infty$, such that $0 < \liminf \frac{K}{N} \leq \limsup \frac{K}{N} < \infty$. Then, for $n = 0, 1, 2, \dots$,

$$\frac{1}{N} \text{tr} \mathbf{D} \mathbf{B}^n - \frac{(-1)^n}{n!} \frac{1}{N} \text{tr} \mathbf{D} \mathbf{T}_n \xrightarrow{\text{a.s.}} 0$$

where \mathbf{T}_n is given by Theorem 2. This implies in particular,

$$\mu_n - \bar{\mu}_n \xrightarrow{\text{a.s.}} 0.$$

Loosely speaking, Theorem 1 states that, for large matrix dimensions, the e.s.d. $F^{\mathbf{B}}$ of the matrix \mathbf{B} can be closely approximated by a deterministic distribution function F . Thus, the optimal weighting vector \mathbf{w} can be approximated by replacing the moments μ_n of $F^{\mathbf{B}}$ in (8) by the moments $\bar{\mu}_n$ of F . Using the result of Theorem 2, we can compute an approximate weight vector $\bar{\mathbf{w}} = [\bar{w}_0 \dots \bar{w}_{L-1}]$ as

$$\bar{\mathbf{w}} = \bar{\Phi}^{-1} \bar{\varphi} \quad (10)$$

where $\bar{\Phi} \in \mathbb{R}_+^{L \times L}$ and $\bar{\varphi} \in \mathbb{R}_+^L$ are defined by

$$\begin{aligned} [\bar{\Phi}]_{ij} &= \bar{\mu}_{i+j} + \sigma^2 \bar{\mu}_{i+j-1} \\ [\bar{\varphi}]_i &= \bar{\mu}_i. \end{aligned} \quad (11)$$

IV. ASYMPTOTIC PERFORMANCE ANALYSIS

We consider now the asymptotic performance of the polynomial expansion receiver in terms of the received SINR γ_k for a given transmitter k . With weight vector \mathbf{w} , the k th element \hat{x}_k of the estimated vector $\hat{\mathbf{x}}$ reads

$$\hat{x}_k = \mathbf{h}_k^H \sum_{l=0}^{L-1} w_l \mathbf{B}^l (\mathbf{H} \mathbf{x} + \mathbf{n}). \quad (12)$$

One can easily show that the associated SINR γ_k can be expressed as [6, Eq. (18)]

$$\gamma_k = \frac{\mathbf{w}^T \boldsymbol{\varphi}_k \boldsymbol{\varphi}_k^T \mathbf{w}}{\mathbf{w}^T (\boldsymbol{\Phi}_k - \boldsymbol{\varphi}_k \boldsymbol{\varphi}_k^T) \mathbf{w}} \quad (13)$$

where $\boldsymbol{\Phi}_k \in \mathbb{R}_+^{L \times L}$ and $\boldsymbol{\varphi}_k \in \mathbb{R}_+^L$ are given as

$$\begin{aligned} [\boldsymbol{\Phi}_k]_{ij} &= [\mathbf{B}^{i+j}]_{kk} + \sigma^2 [\mathbf{B}^{i+j-1}]_{kk} \\ [\boldsymbol{\varphi}_k]_i &= [\mathbf{B}^i]_{kk}. \end{aligned} \quad (14)$$

The next theorem provides a tight deterministic approximation of the terms $[\mathbf{B}^n]_{kk} = \mathbf{h}_k^H \mathbf{B}^{n-1} \mathbf{h}_k$ in the asymptotic limit.

Theorem 4: Under the assumptions of Theorem 3, the following convergence holds:

$$[\mathbf{B}^n]_{kk} - \bar{\mu}_n^k \xrightarrow{\text{a.s.}} 0$$

where

$$\bar{\mu}_n^k = \sum_{i=0}^{n-1} \bar{\mu}_{n-i-1}^k \frac{(-1)^i}{i!} \frac{1}{K} \text{tr} \mathbf{R}_k \mathbf{R}_k^H \mathbf{T}_i, \quad n \geq 1$$

and \mathbf{T}_n is given by Theorem 2. The initial values of the recursion are $\bar{\mu}_0^k = 1$ and $\mathbf{T}_0 = \mathbf{I}_N$.

Proof of Theorem 4: The proof follows the same steps as [6, Theorem 1] and will not be given here. ■

Replacing $[\mathbf{B}^n]_{kk}$ in (14) by $\bar{\mu}_n^k$ and \mathbf{w} in (13) by $\bar{\mathbf{w}}$, we can obtain a deterministic approximation of the SINR γ_k at the output of the polynomial expansion receiver.

V. NUMERICAL RESULTS

Consider a MAC from $K = 40$ single-antenna transmitters to a receiver with $N = 100$ antennas. We use an extended version of Jake's model [4] for the generation of the matrices \mathbf{R}_j . Let $\mathbf{R}_j = \Theta_j^{1/2}$ and $\Theta_j \in \mathbb{C}^{N \times N}$ be defined as

$$[\Theta_j]_{kl} = \frac{1}{\phi_{\max}^j - \phi_{\min}^j} \int_{\phi_{\min}^j}^{\phi_{\max}^j} \exp\left(\frac{2\pi i}{\lambda} d_{kl} \cos(x)\right) dx$$

where $d_{kl} = 2\lambda(k-l)$ and $\phi_{\min}^j, \phi_{\max}^j$ are drawn independently from the intervals $[-\pi, 0]$ and $[0, \pi]$, respectively. The interval $[\phi_{\min}^j, \phi_{\max}^j]$ can be seen as the angular spread of the signal from transmitter j , λ is the wave length, and d_{kl} is the spacing between the receive antennas k and l . We assume Rayleigh fading channels, i.e., \mathbf{w}_j in (2) are independent standard complex Gaussian vectors. The covariance matrices Θ_j are chosen at random at the beginning and then kept fixed while we average over many realizations of the channel matrix \mathbf{H} . We denote by $\text{SNR} = 1/\sigma^2$ the transmit signal-to-noise ratio.

Fig. 1 shows the average received SINR $\mathbb{E}[\gamma_k]$ of a randomly chosen transmitter as a function of the SNR for the matched filter, the LMMSE detector and the polynomial expansion detector with approximate weights for $L = \{2, 3, 6\}$. Markers correspond to simulation results and solid lines to the deterministic SINR approximations. The error bars indicate one standard deviation of γ_k in each direction. Similar to [15], the asymptotic SINR of transmitter k for the LMMSE detector can be easily shown to satisfy

$$\bar{\gamma}_k^{\text{LMMSE}} = \frac{1}{K} \text{tr} \mathbf{R}_k \mathbf{R}_k^H \mathbf{T}(-1/\text{SNR})$$

where $\mathbf{T}(z)$ is given by Theorem 1. We observe a good fit between the deterministic approximations and the simulation results for the average SINR. However, the standard deviation of the SINR increases with L . This is because the higher order moments converge slower to their deterministic approximations and exhibit therefore stronger fluctuations. Nevertheless, the average SINR performance of the polynomial expansion detector with $L = 6$ is already close to the performance of the LMMSE detector.

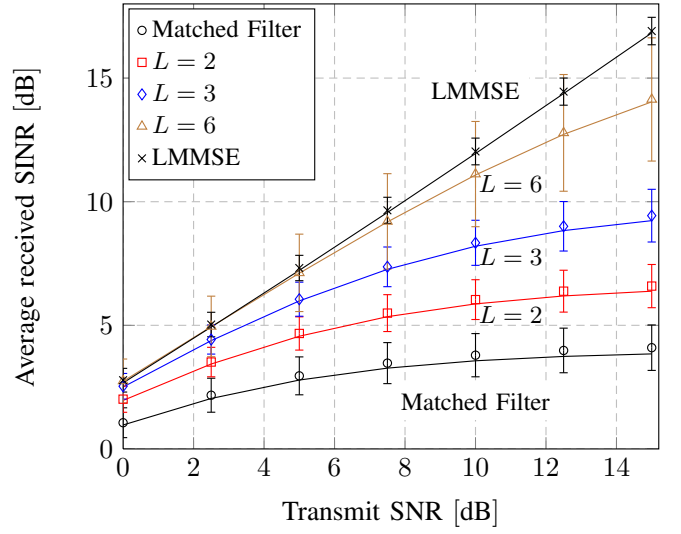


Fig. 1. Average received SINR versus SNR at the output of the matched filter, LMMSE detector and the polynomial expansion detector with approximate weights for different values of L . Markers correspond to simulation results, solid lines to the deterministic SINR approximations. Error bars indicate one standard deviation in each direction.

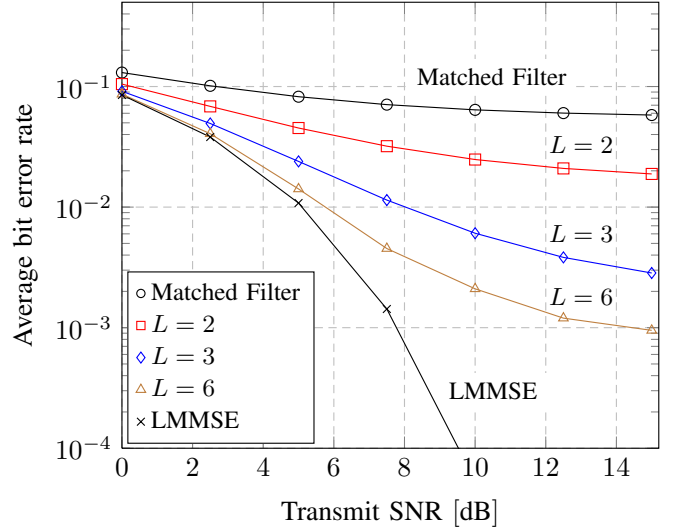


Fig. 2. Average theoretical bit error rate versus SNR for the matched filter, LMMSE detector and the polynomial expansion detector with approximate weights for different values of L .

Fig. 2 depicts the theoretical average bit error rate (BER) over SNR for the different detectors. Assuming binary phase-shift keying (BPSK) modulation and Gaussian interference, the BER is given as $\mathbb{E}[Q(\sqrt{\gamma_k})]$ where $Q(x)$ is the Gaussian tail function. We can clearly see a performance increase of the polynomial expansion detector with L , although the BER saturates at high SNR. Although not explicitly shown here, one can even observe a performance decrease for large values of L . As mentioned before, this is due to the low accuracy of the approximate weights caused by a slow convergence of the higher-order moments to their deterministic approximations.

VI. CONCLUSION

We have derived asymptotically tight deterministic approximations of the moments of a certain class of large random matrices, useful for the study of distributed antenna systems and large antenna arrays. We have applied these moment results to the design of a polynomial expansion detector which significantly reduces the computational complexity of multiuser detection compared to the LMMSE detector. Moreover, we have derived an explicit expression of the asymptotic SINR at the output of this detector and verified its accuracy and performance for finite system dimensions by simulations.

APPENDIX

Outline of the proof of Theorem 2: From Definition 2, it is easy to see that the moments $\bar{\mu}_n$ of the distribution function F can be obtained through successive differentiation of the function $\frac{1}{z}m(-\frac{1}{z})$, i.e.,

$$\begin{aligned}\bar{\mu}_n &= \frac{(-1)^n}{n!} \frac{d^n}{dz^n} \left(\frac{1}{z} m \left(-\frac{1}{z} \right) \right) \Big|_{z=0} \\ &= \frac{(-1)^n}{n!} \int \frac{d^n}{dz^n} \left(\frac{1}{z\lambda + 1} \right) dF(\lambda) \Big|_{z=0} \\ &= \int \lambda^n dF(\lambda).\end{aligned}$$

Consider now the function $\eta(z) = \frac{1}{z}m(-\frac{1}{z})$ for $z \geq 0$ and denote by $\eta_n(z)$ its n th derivative. From Theorem 1, we have

$$\eta(z) = \frac{1}{N} \text{tr} \mathbf{T}_0(z)$$

where

$$\mathbf{T}_0(z) \triangleq \left(z \frac{1}{K} \sum_{j=1}^K \frac{\mathbf{R}_j \mathbf{R}_j^H}{1 + z \delta_{j,0}(z)} + \mathbf{I}_N \right)^{-1}$$

and $(\delta_{1,0}(z), \dots, \delta_{K,0}(z)) \in \mathbb{R}_+^K$ is the unique solution to the K implicit equations:

$$\delta_{j,0}(z) = \frac{1}{K} \text{tr} \mathbf{R}_j \mathbf{R}_j^H \mathbf{T}_0(z), \quad j = 1, \dots, K.$$

Denoting $\mathbf{T}_n(z) = \frac{d^n \mathbf{T}_0(z)}{dz^n}$, we have $\eta_n(z) = \frac{1}{N} \text{tr} \mathbf{T}_n(z)$. Explicit expressions of $\mathbf{T}_n(z)$ can be found by re-writing $\mathbf{T}_0(z)$ with the help of some auxiliary functions and repeated use of the Leibniz-rule for the n th derivative of the product of two functions, i.e., $\frac{d^n (u(x)v(x))}{dx^n} = \sum_{i=0}^n \binom{n}{i} \frac{d^{n-i} u(x)}{dx^{n-i}} \frac{d^i v(x)}{dx^i}$. The resulting set of implicit equations simplifies to a system of recursive equations for $z = 0$. One can easily see that $\mathbf{T}_0(0) = \mathbf{T}_0 = \mathbf{I}_N$ and, consequently, $\delta_{j,0}(0) = \delta_{j,0} = \frac{1}{K} \text{tr} \mathbf{R}_j \mathbf{R}_j^H \forall j$. ■

Outline of the proof of Theorem 3: Both $\frac{1}{N} \text{tr} \mathbf{D} (\mathbf{B} - z \mathbf{I}_N)^{-1}$ and $\frac{1}{N} \text{tr} \mathbf{D} \mathbf{T}(z)$ as defined in Theorem 1 are Stieltjes transforms of finite measures which we denote by π and $\bar{\pi}$. Theorem 1 implies that, almost surely, $\pi - \bar{\pi} \Rightarrow 0$. Similar to the proof of Theorem 2 the moments of π and $\bar{\pi}$ can be respectively expressed as

$$\int \lambda^n \pi(d\lambda) = \frac{1}{N} \text{tr} \mathbf{D} \mathbf{B}^n$$

and

$$\int \lambda^n \bar{\pi}(d\lambda) = \frac{(-1)^n}{n!} \frac{1}{N} \text{tr} \mathbf{D} \mathbf{T}_n.$$

The a.s. weak convergence of π and $\bar{\pi}$ implies by [16, Theorem 25.8 (ii)] that $\int f(\lambda) \pi(d\lambda) - \int f(\lambda) \bar{\pi}(d\lambda) \xrightarrow{\text{a.s.}} 0$, for any bounded continuous function $f(\lambda)$. Relying on [17], one can prove that the support of π is almost surely compact since \mathbf{D} has bounded spectral norm and the spectral norm of \mathbf{B} can be shown to be almost surely bounded. Following similar steps as in [4, Proof of Theorem 2, Part B], one can also show that $\bar{\pi}$ has compact support. Thus, we can relax the assumption of $f(\lambda)$ to be bounded and choose $f(\lambda) = \lambda^n$ to establish the a.s. convergence of the moments of π and $\bar{\pi}$. ■

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